# Multiple Stable Dominance States in the Congo Basin Forests 

John Katembo, Moses B. Libalah, Faustin B Boyemba, Gilles Dauby, Nicolas Barbier

## - To cite this version:

John Katembo, Moses B. Libalah, Faustin B Boyemba, Gilles Dauby, Nicolas Barbier. Multiple Stable Dominance States in the Congo Basin Forests. Forests, 2020, 11 (5), pp.553. 10.3390/f11050553. hal-02586110

## HAL Id: hal-02586110 <br> https://hal.umontpellier.fr/hal-02586110

Submitted on 15 May 2020

HAL is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers. publics ou privés.

Distributed under a Creative Commons Attribution 4.0 International License

Article

# Multiple Stable Dominance States in the Congo Basin Forests 

John M. Katembo ${ }^{1,+(\mathbb{D}}$, Moses B. Libalah ${ }^{2,3, *,+(\mathbb{D}}$, Faustin B. Boyemba ${ }^{1}$, Gilles Dauby ${ }^{4}$ (®D) and Nicolas Barbier ${ }^{4}$ (D)<br>1 Laboratoire d'Ecologie et Aménagement Forestier (LECAFOR), Université de Kisangani, B.P. 2012 Kisangani, Democratic Republic of the Congo; jmukirania1@yahoo.fr (J.M.K.); faustinboyemba@yahoo.fr (F.B.B.)<br>2 Department of Plant Biology, Faculty of Science, University of Yaoundé I, P.O. BOX 812 Yaoundé, Cameroon<br>3 Plant Systematics and Ecology Laboratory (LaBosystE), Higher Teachers' Training College, Yaoundé, University of Yaoundé I, P.O. Box 047 Yaoundé, Cameroon<br>4 AMAP, Univ. Montpellier, IRD, CIRAD, CNRS, INRAE, 34394 Montpellier, France; gildauby@gmail.com (G.D.); nicolas.barbier@ird.fr (N.B.)<br>* Correspondence: libalah_moses@yahoo.com; Tel.: +237-6-7729-4617<br>$\dagger$ These authors contributed equally.

Received: 27 March 2020; Accepted: 12 May 2020; Published: 14 May 2020


#### Abstract

Understanding the dynamics of dominant tree species in tropical forests is important both for biodiversity and carbon-related issues. We focus on the Congo Basin (East of Kisangani) to investigate the respective roles of topographic/soil gradients and endogenous dynamics in shaping local variations in dominance. We used a dataset of 301 -ha plots, in which all trees above 10 cm diameter at breast height (DBH) were censused. Soil samples were analyzed for standard pedologic variables and a digital elevation model permitted to infer topography and hydromorphy. We found that this forest is characterized by variations in the abundance of three dominant species: Petersianthus macrocarpus (P.Beauv.) Liben (PM), Gilbertiodendron dewevrei (De Wild.) J.Leonard (GD) and Julbernardia seretii (De Wild.) Troupin (JS). These variations occur independently of substratum or topography variations. At plot level, the local relative abundance never reached $50 \%$, the threshold for monodominance proposed in the literature, but rather progressively increased to reach higher values for canopy trees ( $>60 \mathrm{~cm} \mathrm{DBH}$ ), where this threshold could be exceeded. We found no sign of shifting compositional dynamics, whereby the dominant species would switch between the canopy and the undergrowth. Our results, therefore, support the possibility of the existence of stable dominance states, induced by endogenous processes, such as biological positive feedbacks fostering monodominance. We also document a strong relation between monodominance level and alpha diversity, when giving more weight to abundant species which is expected $\left(R^{2}=0.79\right)$ but also when giving more weight to rare species $\left(R^{2}=0.37\right)$, showing that monodominance influences tree species richness patterns. Structural differences existed between groups, with the PM group having more (and on average smaller) stems and lighter wood on average, but paradoxically also higher biomass and basal area.


Keywords: DR Congo; forest dynamics; Gilbertiodendron dewerei; Julbernadia seretii; monodominance; soil; succession; Petersianthus macrocarpus

## 1. Introduction

The study of tree assemblages in tropical forests is gaining new impetus with the need to assess carbon emissions at high precision and resolution, while limiting the erosion of diversity and promoting sustainable forest management [1-3]. Heterogeneity in forest composition, diversity and structure is present at all spatial scales. Broad biogeographic variations can be linked to
speciation, extinction and migration events [4-6], climatic gradients, and the influence of specific geomorphological contexts inducing extreme edaphic conditions (e.g., hydromorphy or flooding, poor sandy soils, salinity, etc.) [7-11]. Identifying the drivers of forest heterogeneity at local (landscape) scales, over milder environmental gradients, requires careful analysis. The ecological literature is replete with possible mechanisms of deterministic or stochastic, endogenous or exogenous nature, leading to more or less stable variations in diversity and composition. Patches of contrasted composition could be induced by chance events leading to a random drift driven by stochastic mortality and recruitment [4], by deterministic processes such as biological feedbacks (e.g., density dependent dynamics) and interactions, or simply by the blueprint of soil and topographic variations acting as a filter on the adapted species. Time can also play a major role, via perturbations and succession, which can lead to a locally patchy, but overall homogeneous landscape [12]. These different mechanisms need not be happening independently of each other. For instance, biological feedback along successional gradients will alter both edaphic and (micro-) climatic conditions, leading in some cases to similar climax vegetation from different initial edaphic conditions [13]. Given this overwhelming complexity, how can we hope to glean the necessary pieces of understanding that could help in orienting management or conservation? Can a portion of the heterogeneity in lowland terra firme forests, outside blatantly extreme edaphic conditions, be considered as relatively stable and unambiguously driven either by exogenous (e.g., edaphic) or endogenous (e.g., biotic) processes, or is it all the result of chance dispersion or perturbation events and hence perpetually drifting [4]?

The case of monodominant tropical forests has been the focus of sustained research efforts over the past few decades, because they stand as a particularly prominent feature among otherwise very diverse forests. The quantitative definition of monodominance may be quite subjective: a $50 \%$ relative abundance threshold for trees $>10 \mathrm{~cm}$ diameter at 1.3 m breast height ( DBH ) has been proposed [14,15], but a recent meta-analysis in the Amazon region suggests this threshold might be too restrictive [16]. Dominance is obviously a continuous measure, but setting thresholds is useful for inter-comparisons and classification purposes. In any case, focusing on monodominant forests might help understanding crucial processes in more diverse forests. Questions indeed immediately arise regarding the different mechanisms listed earlier. Are these patches located in peculiar topographic/edaphic positions? Among the different types of monodominance described by [14], does the dominant species remain persistent for numerous tree generations (type I) or is the monodominant state an early successional category (Type II) or the result of mostly stochastic demographic and limited dispersion events? What biological feedback mechanisms might be at play? Mycorrhizae or allelopathy are often cited [17]. What about the apparent violation of the negative density-dependence [18,19]? Are monodominant patches expanding over surrounding forests, or is the mosaic stabilized by some feedback mechanisms?

A particularly studied case is that of widespread monodominance occurring across central African forests, with species like Gilbertiodendron dewevrei (De Wild.) J. Leonard or Julbernardia seretii (De Wild.) Troupin (Fabaceae-Caesalpinioideae) covering sometimes very extensive, although patchy extends across the Congo Basin and its surroundings [20,21]. According to [22] no less than two-thirds of the Northeastern part of the Congo Basin forest would in fact be covered by monodominant forests. Soil chemical fertility has not been found to be significantly different from surrounding forest types [23-25] apart from organic matter content [25], and geomorphological (catena) positions can also be considered to be similar [15,26,27]. Although [28] showed that some tree species were sensitive to edaphic gradients (from sandy to clayey soil) near Kisangani (DRC), these gradients did not seem to affect dominant species. In another study in Cameroon, ([29] in review) the authors also noticed that soil effects tend to influence relatively rare species, but not the most common ones. Therefore, the role of soil-related properties on monodominance seem little supported outside of extreme environments (swamps, white sands, etc.).

The general question addressed here pertains to the possible endogenous nature of forest dominance patterns, the stability of the different states, and their consequences in terms of biomass (carbon stocks) and biodiversity. In the absence of long term sylvicultural experiments, only macroecological (correlative)
approaches can offer clues about processes driving monodominance in the Congo basin. The forests around Kisangani (DRC) present contrasted compositional and structural patterns hosting interlaced patches of mixed-species forest stands and monodominant stands of G. dewevrei and/or J. seretii [30]. The soils are also heterogeneous both at the landscape and at fine-scales [25,28]. We established 30 one-hectare tree plots with the aims to identify: (1) floristic groups characterized by dominant species, (2) local environmental drivers (soil, topography and forest stand attributes) potentially affecting floristic composition and (3) trends in composition across diameter size classes. We tested the hypothesis that variation in the local abundance of different dominant species is independent of environmental drivers (i.e., classical monodominance, sensu [17]) but possibly a signal of successional stages. Parallel to the question of drivers of local compositional dominance, we also document variations in forest stand attributes (stem density, basal area, above-ground biomass, the quadratic mean diameter and community wood density) and alpha diversity between the compositional groups in comparison with peculiarities noted by previous studies [31-33].

## 2. Material and Methods

### 2.1. Study Site

Data for this study were collected in a lowland moist dense tropical forest in the northeast of the Central Congo Basin, referred to as the Cuvette Centrale. Shown in Figure 1, the site is situated 79 km East of Kisangani between $0^{\circ} 25^{\prime} 34^{\prime \prime}-0^{\circ} 34^{\prime} 51^{\prime \prime} \mathrm{N} ; 25^{\circ} 46^{\prime} 25^{\prime \prime}-25^{\circ} 53^{\prime} 10^{\prime \prime}$ E, covering an area of $260 \mathrm{~km}^{2}$. The forest is mostly semideciduous but evergreen monodominant stands of Gilbertiodendron dewevrei and Julbernadia seretii are common [30]. The climate is of the tropical rainforest type, Af in Köppen [34,35]. Two rainy seasons are distinguished with mean monthly precipitation above 100 mm and an annual average of 1772 mm and one dry season (December to February) with monthly precipitation lower than $100 \mathrm{~mm}[36,37]$. Average daily temperature is $25.2^{\circ} \mathrm{C}$. Ferralitic soils dominate nonhydromorphic areas [25] and the topography is characterized by rocky outcrops (inselbergs) and sandy plateaus in the South. The local hydrographic network is led by the Selesele River which collects surrounding waters from the south and flows into the Tshopo River which is the largest in the study area. The latter cuts across the area in its northernmost part and is responsible for large alluvial or clay deposits.


Figure 1. Distribution of one-hectare plots along an altitudinal gradient. Dots represent 1-ha plots; dot sizes are proportional to alpha diversity and colors represent plots within the same $k$-means derived groups.

### 2.2. Plot Establishment and Tree Census

We established thirty noncontiguous $100 \mathrm{~m} \times 100 \mathrm{~m}$ (1-ha plot) distributed around the study area along altitudinal transects, as shown in Figure 1. To do this, we first surveyed a 100 m square area by placing stakes at the four corners of the intended area. Subsequently, we divided the 100 m square area into 25 quadrats of 20 m square each while avoiding destruction of all woody plants, using a Suunto ${ }^{\circledR}$ KB14 compass (Suunto ${ }^{\circledR}$, Vantaa, Finland) and measuring tapes. A consumer grade Garmin ${ }^{\circledR} 62$ GNSS device (Garmin ${ }^{\circledR}$, Olathe, KS, USA) was used to geolocate each 1-ha plot, by recording positions every 20 m around the plot contour to attenuate the effect of signal scattering in the vegetation [38]. We then censused all trees by the following methods: (i) the diameter of all woody stems (excluding lianas) $\geq 10 \mathrm{~cm}$ at 1.3 m breast height (DBH) or above buttresses and deformations was measured using a diameter tape. (ii) Maximum height of at least 40 trees within each 1-ha plot across the full DBH range was measured using a Trimble ${ }^{\circledR}$ LaserAce ${ }^{\text {TM }} 1000$ rangefinder (Trimble ${ }^{\circledR}$, Westminster, CO, USA). The sine method was used to measure maximum tree height [39], which is done by standing at a given distance from the tree and open canopy, using the LaserAce ${ }^{\mathrm{TM}} 1000$ rangefinder to measure the distance to the top and base of the tree and the angles from horizontal. (iii) We identified all trees to species level based on our field experience but also collected botanical samples for herbarium verification. All voucher specimens are stored at the University of Kisangani and the herbarium of the Université Libre de Bruxelles (BRLU). Nomenclature of plant families follows the Angiosperm Phylogeny Groups IV [40] and species names are according to the 2016 African Plants Database (version 3.4.0) [41].

### 2.3. Soil Data

Four soil samples were collected along the central diagonal of each plot. We focused soil collection on the upper $0-30 \mathrm{~cm}$ layer (excluding litter) to make comparisons with other local and regional studies ([29] in review). We collected 60 g from each point and 120 soil samples in total from the 30 1-ha plots. Each sample was numbered, separately air-dried and sieved with a 2 mm sieve prior to analyzing soil textural properties and a 0.5 mm sieve prior to analyzing chemical properties of the soil. The percentage proportion of clay ( $\%,<2 \mu \mathrm{~m}$ ), silt ( $\%, 2-50 \mu \mathrm{~m}$ ) and sand ( $\%, 50 \mu \mathrm{~m}-2 \mathrm{~mm}$ ) were obtained by pipette sedimentation column. Soil pH was measured by 1:2.5 soil:deionized water ratio. Total organic carbon ( $\mathrm{Cg} / 100$ ) and total nitrogen $(\mathrm{Ng} / 100)$ were first analyzed by dry combustion and measured using a thermal conductive detector. Available phosphorus ( $\mathrm{P} \mathrm{mg} / \mathrm{kg}$ ) was extracted at pH 8.5 using 0.5 M ammonium fluoride and 0.5 M sodium bicarbonate via the spectrometric method. The concentration of calcium ( $\mathrm{Camg} / \mathrm{kg}$ ), magnesium $(\mathrm{Mg} \mathrm{mg} / \mathrm{kg})$ and potassium ( $\mathrm{K} \mathrm{mg} / \mathrm{kg}$ ) were extracted during saturation using 1 M ammonium acetate and the residues were washed with ethyl alcohol. Soil analysis was conducted in the Laboratoire pédologique de l'Institut Facultaire des Sciences Agronomiques de Yangambi (IFA-Yangambi, DR Congo).

### 2.4. Topographic Data

We used the centroids of the 20 GPS coordinates of each 1-ha plot (see Section 2.2) to extract four variables proxies of local topography and hydromorphy, namely, the height above nearest drainage (HAND) index [42], topographic moisture index (TMI) [43], plot altitude, flow accumulation and slope. Data for these variables were at 1-arc second ( $30-\mathrm{m}$ ) resolution from Shuttle Radar Topography Mission (SRTM). This was obtainable at USGS-EarthExplorer (https://earthexplorer.usgs.gov/) by means of the routines implemented in RSAGA [44]. The same data were used to visualize the distribution of the plots.

### 2.5. Forest Stand Attributes

Five parameters describing forest attributes were calculated, including maximum tree height, basal area, stem density, wood density and quadratic diameter. We used the randomly selected trees heights measured from the 1-ha plots ( 40 to 191 trees measured per plot; Table S1) to predict the heights of unmeasured trees. To do this, we calibrated, for each plot, an allometric model of the Weibull form (Equation (1)) [45], using the modelHD function in the BIOMASS R-package [46]. We calculated the tree basal area ( $\mathrm{G}, \mathrm{m}^{2}$, Equation (2)) as the sum of individual tree sections (diameter) at breast height in each plot. Stem density ( N , counts), quadratic mean diameter ( Dg , cm, Equation (3)), mean height ( $\mathrm{H}, \mathrm{m}$ ) and mean wood density weighted basal area (WD, $\mathrm{g} \mathrm{cm}^{-3}$ ) were also derived. We employed Equation (4) of [47] to predict tree above-ground biomass (AGB, Mg ha- ${ }^{-1}$; Equation (4)) based on individual's DBH, tree height (H) and species mean wood density (WD) from the global repository (DRYAD; [48]).

$$
\begin{gather*}
H=a \times\left(1-\exp \left(-(D B H / b)^{C}\right)\right)  \tag{1}\\
G=\sum_{i=1}^{N} \frac{\pi \times D B H}{4}  \tag{2}\\
D_{g}=\sqrt{\frac{\sum D B H_{i}^{2}}{N}}  \tag{3}\\
A G B=0.0673 \times(W D \times H \times D B H)^{0.976} \tag{4}
\end{gather*}
$$

where $a, b$ and $c$ are parameter estimates for Equation (1).

### 2.6. Multivariate Analysis

We identified the major floristic groups by means of a non-symmetric correspondence analysis (NSCA; [49-51]). This method is suitable here because it emphasizes abundant species from the plot profile via a 281 species by 301 -ha plots contingency table. We focused only on fully and correctly identified species ( $98 \%$ ) and excluded those with possible taxonomic issues or identified only at the genus level ( $<2 \%$ of all individuals). We extracted plots' and species' scores from the NSCA and grouped plots into three clusters based on a $k$-means clustering.

We determined correlates of the major floristic groups via the environmental vector fitting approach using the envfit function of the vegan R-package [52]. This non-constrained analysis simply displays the projection of linear correlations between variables of interest (potential drivers or forest structure variables) and floristic gradients in the chosen plane. The length and direction of the arrows of the multivariate environmental fitting are indicative of the strength and sign of the correlations, respectively. We then used a permutation test to determine the significance of each relationship. The latter analysis was conducted separately for the relationships between floristic gradients and forest attributes on the one side and floristic gradients and soil, topography-related variables on the other.

### 2.7. Characterization of Taxonomic Alpha Diversity

We compared taxonomic diversity at plot level (alpha diversity) across levels of monodominance (i.e., frequency of the most frequent species) and floristic groups. Taxonomic diversity within communities is influenced by the number of sampled individuals ( N ), the species richness ( S ), species relative abundance (evenness) and spatial aggregation [53]. Given these components of diversity change, it is necessary to use several diversity metrics [53,54]. Here, we were interested in changes in species abundance distribution which are influenced by both $S$ and evenness. We used an individual-based rarefaction (IBR) framework transformed into equivalent number of species (ENS) [55]. IBR consists in calculating the expected number of species given $k$ randomly sampled individuals $\left(S_{k}\right)$. Increasing the parameter $k$ proportionally increases the weight given to rare species. $S_{k=2}$ corresponds to the Gini-Simpson index which is an index of the diversity of abundant species while $S_{k=300}$ gives much more weight to rare species and can thus be considered as an index of species
richness unbiased by sample size [56]. The advantage to transform IBR into ENS ( $S_{k}$ converted into $E N S_{k}$ ) is that interpretation of diversity expressed as equivalent number of species values is more intuitive $[55,57]$. Finally, visualizing how $E N S_{k}$ varies with $k$ indicates the level of evenness of species distribution within the community: if all species are equally frequent, $E N S_{k}$ remains constant across $k$ values, whereas in a community dominated by one species while all other species are rare, will have $E N S_{k}$ increasing steeply with $k$. We computed $E N S_{k}$ in the R environment using functions developed by [55] and available in github (github.com/gdauby/hurlbert_ens).

We expect to observe a very good correlation between $E N S_{k=2}$ and level of monodominance because high monodominance level means low diversity of dominant species. For higher $k$ values, a strong correlation is expected only if strong monodominance also implies low diversity of rare species.

### 2.8. Insight into the Forest Dynamics of the Study Area

We used a synchronous approach to investigate possible compositional shifts or trends in the forest floristic composition. Trees of each species were split into two size categories based on a cut-off at 60 cm DBH (i.e., $10-60 \mathrm{~cm}$ and $>60 \mathrm{~cm}$ ). This allowed to distinguish the 'mature' canopy while keeping sufficient sample sizes, i.e., around 25 trees (range: 17-36) in the canopy [58].

## 3. Results

### 3.1. Species Compositional Gradients and Stand Attributes

A total of 13433 trees ( $\geq 10 \mathrm{~cm}$ at 1.3 m diameter breast height (DBH)) were recorded across the 30 one-ha plots belonging to 281 species, 179 genera and 44 families. The frequency and relative abundance in the plots were dominated by a few prominent species such as Julbernardia seretii (De Wild.) Troupin, Gilbertiodendron dewevrei (De Wild.) J.Leonard (both Leguminosae), Petersianthus macrocarpus (P.Beauv.) Liben (Lecythidaceae), Greenwayodendron suaveolens (Engl. \& Diels) Verdc. and Anonidium mannii (Oliv.) Engl. \& Diels (both Annonaceae), as shown in Figure 2. The first two species (J. seretii and G. dewevrei) were present in about half of the plots scoring relatively high mean abundance ( $5.6 \%$ and $7.6 \%$ respectively; Figure 2A) but G. dewevrei scored the highest maximum abundance of $39.1 \%$, as shown in Figure 2B. Other species, namely P. macrocarpus, G. suaveolens and A. mannii were more frequent, i.e., present in all or nearly all plots and also abundant. Each of these species scored mean percentage abundance of 5.8, 4.6 and $4.9 \%$, respectively, as displayed in Figure 2A, and up to $23.9 \%$ maximum local relative abundance for $A$. mannii, as displayed in Figure 2B.

Four species (J. seretii, G. dewevrei, P. macrocarpus and A. mannii) characterized three major floristic groups as depicted by the $k$-means clustering and non-symmetric correspondence analysis (NSCA), as shown in Figure 3A. The first two compositional axes of the NSCA together constituted $48.25 \%$ of the total floristic variance. The first axis (negative side) was clearly pulled by G. dewevrei while the second axis opposed plots dominated by P. macrocarpus (positive side) and J. seretii and A. mannii on the negative side.

Changes in forest stand attributes were substantial with wide range values recorded for each of the 30 1-ha plots; number of stems (332-557; mean: 447.8 count), quadratic mean diameter (27.2-35.5; mean: 30.2 cm ), basal area (22.7-39.5; mean: $31.9 \mathrm{~m}^{2}$ ) and above-ground biomass (331.9-590.6 mean: $447.8 \mathrm{Mg} \cdot \mathrm{ha}^{-1}$ ). The average maximum tree height per plot ranges between 16.9 m and 28.4 m , shown in Table S1.


Figure 2. Mean and maximum local relative abundance against species frequency of occurrence in the 301 -ha plots. (A) Mean local relative abundance (only accounting for plots where the species is present) and (B) maximum local relative abundance. Only species frequencies $>18$, mean abundance $>4.5$ and max abundance $>10$ are shown for legibility.


Figure 3. Linking floristic gradients to environmental variables. (A) Biplot of first two axes of floristic composition depicted from nonsymmetrical correspondence analysis (NSCA); dots represent 1-ha plots and dot colors represent groups from $k$-means clustering. (B) Projected correlation with soil and topographic variables. (C) Projected correlation with forest attributes. Grey arrows represent nonsignificant correlations and bold arrows represent significant correlations after permutation test. All variables are defined in Table 1.

### 3.2. Relationships with Environmental Variables and Forest Attributes

The relationships were generally weak between the main floristic axes (NSCA 1 and NSCA 2) and topographic, soil variables shown in Figure 3B and forest attributes shown in Figure 3C. The correlations between soil and topography variables and floristic gradients were not significant according to the permutation tests, as presented in Table 1. Nevertheless, 12 out of 13 plots of J. seretii occupied sites on higher altitude and height above drainage (HAND), as shown by a post hoc test (multiple comparison test with Bonferonni adjusted thresholds, Table 2) and the map of the floristic groups to altitudinal gradient (Figure 1), possibly with poorer sandy soils, as displayed in Figure 3B. G. dewevrei stands were apparently (but not significantly) correlated with slopes and showed no preference for lower topographic or hydromorphic positions. P. macrocarpus stands seemingly favored richer, lowland fine textured soils. Two forest stand attributes significantly correlated with the floristic composition axes according to the permutation test, as shown in Figure 3C and Table 1. Plot level wood density (WD) was negatively correlated with the first NSCA axis, due to the relatively high wood density of G. dewevrei $\left(0.707 \mathrm{~g} \mathrm{~cm}^{-3}\right)$. Plot level stand density ( N ) was strongly positively correlated with the second NSCA axis, indicating high stem numbers in P. macrocarpus forests, as shown in Table S1. Basal area (G), quadratic mean diameter ( Dg ) and above-ground biomass (AGB) were not significantly correlated with the second NSCA axis, suggesting higher biomass in P. macrocarpus forests, despite larger trees on average ( Dg ) in the two other forest types (Table S1).

Table 1. Correlation between the first two non-symmetric correspondence analysis (NSCA) axes of species composition and soil, topographic variables and stand structure variables. Significance ( $P$-value) provided by a permutation test (999 permutations) and bold-faced characters indicate statistically significant correlation at $P<0.05$. HAND: height above nearest drainage; TMI: topographic moisture index; FlowAcc: flow accumulation; G: tree basal area; N: stem density; Dg: quadratic mean diameter; H: mean tree height; AGB: tree above-ground biomass; WD: species mean wood density.

|  | Range | Axis1 | Axis2 | $P$-Value | $R^{2}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Soil |  |  |  |  |  |
| Sand | 33-90 | 0.349 | -0.937 | 0.534 | 0.046 |
| Clay | 6-52 | -0.524 | 0.852 | 0.530 | 0.045 |
| Silt | 4-15 | 0.307 | 0.952 | 0.495 | 0.056 |
| Nitrogen | 5.39-16.73 | -0.728 | 0.686 | 0.259 | 0.092 |
| Carbon | 6.388-7.063 | -0.074 | 0.997 | 0.536 | 0.045 |
| pH | 3.9-4.162 | 0.747 | 0.664 | 0.765 | 0.020 |
| Phosphorus | 0.035-0.415 | -0.985 | 0.172 | 0.816 | 0.011 |
| Magnesium | 0.22-0.35 | -0.775 | 0.632 | 0.926 | 0.006 |
| Calcium | 0.083-0.145 | -0.525 | 0.851 | 0.806 | 0.017 |
| Potassium | 1.7-3.825 | 0.351 | -0.936 | 0.611 | 0.037 |
| Topography |  |  |  |  |  |
| Altitude | 461-687 | 0.076 | -0.997 | 0.322 | 0.079 |
| Slope | 2.64-12.63 | -0.837 | -0.547 | 0.432 | 0.056 |
| HAND | -451-247 | 0.020 | -1.000 | 0.500 | 0.049 |
| TMI | -3.35-2.9 | 0.985 | -0.172 | 0.473 | 0.059 |
| FlowAcc | -6-1.185 | 0.803 | 0.596 | 0.181 | 0.124 |
| Stand structure |  |  |  |  |  |
| G | 24.22-39.58 | 0.381 | 0.925 | 0.103 | 0.159 |
| N | 322-557 | 0.334 | 0.942 | 0.001 | 0.466 |
| Dg | 27.26-35.51 | -0.307 | -0.952 | 0.233 | 0.102 |
| H | 18.25-32.37 | 0.186 | 0.983 | 0.496 | 0.048 |
| AGB | 333.43-590.69 | 0.112 | 0.994 | 0.568 | 0.044 |
| WD | 0.56-0.68 | -0.983 | -0.186 | 0.018 | 0.235 |

Table 2. Relationship between floristic groups and structural and topo-soil-related variables. Significant differences at $P<0.05$ for each variable between the floristic groups, obtained from multiple comparison tests with Bonferroni adjustment, are shown between brackets. For a given variable, groups with the same bracketed letters are not significantly different.

| Variables | GROUP1 $(n=5)$ | GROUP2 $(n=13)$ | GROUP3 $(n=12)$ |
| :---: | :---: | :---: | :---: |
|  | Soil |  |  |
| Clay | $23.2 \pm 17.31(a)$ | $21.92 \pm 15.1(a)$ | $32.75 \pm 11.67(a)$ |
| Sand | $69.6 \pm 20.13(a)$ | $69.62 \pm 18.9(a)$ | $58.25 \pm 11.44(a)$ |
| Silt | $7.2 \pm 3.11(a)$ | $8.46 \pm 4.25(a)$ | $9 \pm 3.72(a)$ |
| Nitrogen | $8.24 \pm 2.16(a)$ | $9.82 \pm 2.97(a)$ | $8.67 \pm 1.62(a)$ |
| Carbon | $6.77 \pm 0.15(a)$ | $6.75 \pm 0.17(a)$ | $6.66 \pm 0.16(a)$ |
| pH | $4 \pm 0.08(a b)$ | $3.97 \pm 0.06(b)$ | $4.06 \pm 0.06(a)$ |
| Phosphorus | $0.08 \pm 0.02(a)$ | $0.12 \pm 0.11(a)$ | $0.1 \pm 0.05(a)$ |
| Magnesium | $0.29 \pm 0.02(a)$ | $0.3 \pm 0.03(a)$ | $0.28 \pm 0.03(a)$ |
| Calcium | $0.11 \pm 0.003(a)$ | $0.12 \pm 0.01(a)$ | $0.11 \pm 0.02(a)$ |
| Potassium | $2.21 \pm 0.36(a)$ | $2.71 \pm 0.66(a)$ | $2.41 \pm 0.46(a)$ |
| Topography |  |  |  |
| Altitude | $513.2 \pm 27.24(b)$ | $585.85 \pm 58.94(a)$ | $505.58 \pm 30.8(b)$ |
| Slope | $5.32 \pm 1.24(a)$ | $5.92 \pm 2.95(a)$ | $5.19 \pm 1.97(a)$ |
| HAND | $-415.2 \pm 21.91(b)$ | $-347.38 \pm 53.18(a)$ | $-415.33 \pm 22.69(b)$ |
| TMI | $-1.96 \pm 2.16(a)$ | $-1.89 \pm 2.15(a)$ | $0.14 \pm 2.31(a)$ |
| FlowAcc | $-3.5 \pm 3.42(a)$ | $-4.37 \pm 3.12(a)$ | $-3.8 \pm 3.28(a)$ |
| Stand structure |  |  |  |
| G | $31 \pm 2.33(a)$ | $30.73 \pm 5.11(a)$ | $33.77 \pm 3.99(a)$ |
| N | $418 \pm 40(b)$ | $417 \pm 58(b)$ | $493 \pm 56(a)$ |
| Dg | $30.79 \pm 1.95(a)$ | $30.66 \pm 2.5(a)$ | $29.56 \pm 1.55(a)$ |
| AGB | $453.17 \pm 58.83(a)$ | $442.23 \pm 79.33(a)$ | $461.29 \pm 60.06(a)$ |
| H | $25.64 \pm 5.69(a)$ | $27.21 \pm 4.41(a)$ | $26.82 \pm 2.9(a)$ |
| WD | $0.67 \pm 0.01(a)$ | $0.63 \pm 0.03(b)$ | $0.63 \pm 0.01(b)$ |
|  |  |  |  |

### 3.3. Alpha Diversity across Monodominance Levels and Floristic Groups

As expected, we observed a strong relationship between the level of monodominance and alpha diversity expressed as $E N S_{k=2}$ which essentially estimates the diversity of abundant species ( $R^{2}$ of the linear regression reached 0.79 ). This correlation was still present, although weaker, when giving more weight to rare species ( $R^{2}$ of the linear regression between level of monodominance and $E N S_{k=300}=0.37$ ), as shown in Figure 4. Alpha diversity appeared to vary independently of floristic group regardless of the metric used, as displayed in Figure 4.

### 3.4. Shift in Species Dominance

Overall, the relative abundance of all dominant species in the canopy ( $\geq 60 \mathrm{~cm}$ DBH) was strongly correlated to their own abundance in the lower strata ( $10-60 \mathrm{~cm}$ DBH), and not to the abundance of any other dominant species, as shown in Figure 5. Hence, analyzing differential relative abundance of the three dominant species in the lower strata and in the canopy shown in Figure 6 did not provide evidence of shifts in dominance, in which a species would obviously tend to replace another through time in any of the three floristic groups. We noticed however, that in six plots dominated by J. seretii, the order of dominance in the lower strata was different from that in the canopy, to the benefit of either or both the other dominant species. For the three dominant species, the relative abundance was always lower in the lower strata than in the canopy, as shown in Figure 6. If we focus on cases with plot level dominance exceeding $10 \%$, relative density differences between the canopy and the lower strata were on average $11.3 \%$ for G. dewevrei, $45.6 \%$ for $J$. seretii and $14.1 \%$ for P. macrocarpus. Many species with a relatively high abundance in the lower strata rarely reach the canopy, and hence have a negative density difference, such as $A$. mannii (a typical lower strata tree), with $-21 \%$.


Figure 4. Individual-based rarefaction curves for the 30 plots expressed as equivalent number of species of the expected number of species in $k$ randomly sampled individuals $E N S_{k=2}$. Colors of the curves reflect the level of monodominance expressed as the frequency of the most frequent species. Increase in the $k$ parameter ( x -axis) increases the sensitivity to rare species. The insert on the top left corner shows the relation between alpha diversity and the level of monodominance for $E N S_{k=2}$ thus being insensitive to rare species while the insert on the bottom right corner shows the relation between alpha diversity and the level of monodominance for $E N S_{k=300}$, thus being strongly sensitive to rare species. Colors of the dots in inserts correspond to the three floristic groups identified.


Figure 5. Correlation matrix for size classes (cutoff at 60 cm diameter at breast height (DBH)) of the main dominant species. Cutoffs for diameter class are either $<60 \mathrm{~cm}$ (species name only) or $>60 \mathrm{~cm}$ (species name_60; where tree size exists).


Figure 6. Mean relative abundance of the three main dominant species among floristic groups. Groups are defined from $k$-means clustering on nonsymmetric correspondence analysis scores. (a) In the lower strata ( $\leq 60 \mathrm{~cm} \mathrm{DBH}$ ); (b) in the canopy ( $>60 \mathrm{~cm} \mathrm{DBH}$ ). GD is Gilbertiodendron dewerei; JS is Julbernadia seretii; PM is Petersianthus macrocarpus.

## 4. Discussion

Using a sampling of 30 1-ha plots sampled across a $260 \mathrm{~km}^{2}$ area, we evidenced strong variations in the abundance of dominant species characterized by Gilbertiodendron dewevrei, Julbernadia seretii, Petersianthus macrocarpus and Anonidium mannii, as shown in Figure 3A. The variation in dominance of these species seems to occur nearly independently of substratum variations, as shown in Figure 2B, despite marginally significant differences for the group dominated by J. seretii. This result corroborates previous studies refuting soil characteristics [23-25,59,60] and topography [26] as potential drivers of monodominance in tropical forests. Two forest stand attributes (wood and stem densities) correlated with the floristic gradients, as shown in Figure 3C. Wood density was on average higher in the floristic groups dominated by G. dewevrei. On the contrary, stem density was significantly higher in the P. macrocarpus floristic group. These results taken together could suggest a higher maturity of a late successional stage for the Fabaceae groupings (i.e., G. dewevrei and J. seretii). Contrary to [31], basal area and above-ground biomass (AGB) were not correlated to the dominance by G. dewevrei. Instead, P. macrocarpus stands, despite lighter woods, had a higher basal area and AGB owing to the abundance of both taller and more numerous trees. This counterintuitive result can be related to other studies in Central Africa. For instance, taller trees in denser stands tend to have higher biomass in semideciduous forests compared to evergreen forests [61] or in wet forests compared to moist forests [62,63]. Interestingly, the AGB range in G. dewevrei forests in our study area ( 379 to $533 \mathrm{Mg} \mathrm{ha}^{-1}$, 453 Mg ha $^{-1}$ average, 5 plots) appeared to be slightly lower that the AGB obtained in a similar study in closed canopy forests in Cameroon ( $514 \mathrm{Mg} \mathrm{ha}^{-1}$ average, 23 plots of $G$. dewevrei forests) $[31,32]$. This result may suggest that $G$. dewevrei plots in the study area are still stocking above ground biomass.

Strong correlations were observed between the dominance of a given species in the canopy ( $\geq 60 \mathrm{~cm}$ diameter at 1.3 m breast height (DBH)) and its abundance in the lower strata ( $10-60 \mathrm{~cm} \mathrm{DBH}$ ). This suggests that in this study area the states of dominance in the vegetation are stable across generations, that successional dynamics are very slow or that they are localized to peculiar locations (e.g., forest edges). Dominant species were however less dominant in the lower strata than in the canopy, as shown in Figure 5. Under the hypothesis of stable dynamics, the latter could be explained by biological feedbacks operating after recruitment. The case of P. macrocarpus, considered as a light-demanding nonpioneer species shows that succession to oneself does not necessarily imply shade tolerance. The canopy heterogeneity within one hectare is indeed considerable, leading to a mosaic of patches of contrasted microenvironments [12]. These findings suggest that monodominant forest patches could constitute stable states maintained by endogenous feedbacks [64], such as biological interactions between individuals or species, either positive (mycorrhiza associations, root anastomosis) [31,65,66] or negative (competition for light or nutrients, allelopathy) [14,66,67], short range seed dispersal and
enhanced resistance to pathogens and herbivores leading to mild negative density dependence ([18,19] Hypothesis). An unknown combination of the above mechanisms could lead to stabilizing patches dominated by one or a few species. We cannot, however, discard the hypothesis that widespread historical human disturbance may have been at the origin of state change from one forest type to another, as hypothesized by [68], or that the temporal scale necessary for compositional changes are not properly captured by our synchronic observations. Another possible explanation, involving random drift in the community composition (neutral theory [4]) would imply, at least intuitively, that dominance should be achieved by radically different species across locations, and not repeatedly by the same few species across the landscape or even the continent [29].

Although some of the sampled forests correspond to typical examples of monodominant forests, the $50 \%$ abundance threshold was never reached at plot level. We suggest that the conclusion of recent studies based on this threshold and making inferences about the rarity of monodominant forests (e.g., in the Amazon, [16] and in the Congo Basin, [15]) should be reconsidered. At a $20 \%$ abundance threshold, the same authors reported $18 \%$ of monodominant plots. Indeed, at the $20 \%$ abundance threshold in our study site, $13.3 \%$ of the plots could be considered monodominant, and up to $30 \%$ of the plots had a $15 \%$ dominance level. Moreover, this only considers the dominant species one by one. What perhaps matters more is that the plot composition often tends to become uneven and that most plots are distributed along a gradient of dominance by one or a handful of species.

Unsurprisingly, the level of monodominance strongly influences the diversity of abundant species. This effect is not only the result of stronger unevenness in relative species distribution. Indeed, if monodominance only resulted in changes in evenness, all rarefaction curves in Figure 4 would tend to converge when increasing $k$ values [53]. Monodominance was also, to a lesser extent, correlated to diversity estimated by $E N S_{k=300}$ and thus associated with low richness of both rare and abundant species.

Monodominance is common among species of the Fabaceae [17] and two dominant species (G. dewevrei and J. seretii) in our study site are known to be gregarious and belong to the Fabaceae (Detarioideae). Much has been written on possible mechanisms driving local dominance in this family, such as competitive advantages provided by ectomycorrhizae, shade tolerance or heavy litter accumulation $[65,66]$. However, this focus largely occults the fact that species from other families (in our case Lecytidaceae e.g., P. macrocarpus and Annonaceae e.g., A. mannii) are both more frequent and nearly as locally dominant. Species from other families such as Apocynaceae, Burseraceae and Moraceae [17] have also been reported elsewhere.

Given the importance and prevalence of dominance patterns observed in our site but also elsewhere in the Congo Basin ([29] in review) or in the neotropics [16], and the implications in terms of plant diversity, evolution (e.g., the role of density dependence), forest functioning and structure, we advocate more in-depth studies of the mechanisms leading to the emergence and maintenance of multiple stable states in tropical forests.

## 5. Conclusions

This paper provides new results from tree communities at the heart of the Congo Basin forests documenting the respective roles of environmental gradients and endogenous dynamics in shaping local variations in dominance. We show that a few species compete to dominate tree communities, mostly independently of abiotic (edaphic, topographic) constraints. Moreover, succession seems to be stalling, with the locally dominant species tending to dominate the understory as well. This is evocative of either emergent patterning or multiple stable states induced by biological feedbacks [64]. Although we are not at this stage able to provide a fully documented mechanistic explanation, we think this contribution should help initiate fruitful future interdisciplinary studies involving both experimental biology and mathematical modelling.

We show that the level of monodominance has a significant impact on tree diversity patterns and is thus an important feature to document and understand for efficient biodiversity conservation. Finally,
dominance by one or another species also has implications in terms of forest structure and biomass. The different states are therefore not equivalent regarding carbon stocking or other ecosystem services.

Supplementary Materials: The following are available online at http://www.mdpi.com/1999-4907/11/5/0/s1, Table S1: Summary of topography and forest attributes collected from 30 1-ha plots in the tropical forest in the Uma forest of the Congo Basin.
Author Contributions: Conceptualization, J.M.K., M.B.L., G.D. and N.B.; methodology J.M.K., G.D. and N.B.; formal analysis, J.M.K., M.B.L. and G.D. with help from N.B.; writing-original draft preparation, J.M.K. and M.B.L. with help from N.B., G.D. and F.B.B.; supervision, N.B. and F.B.B.; incorporating critical feedback from all authors, which helped shape the manuscript, N.B., G.D. and F.B.B. All authors have read and agreed to the published version of the manuscript.
Funding: This work was supported by IRD/UMR-AMAP (http://amap.cirad.fr) for soil analyses and two scientific exchange visits (South - South mobility in 2018 and South - North mobility in 2020), CIFOR (Centre for International Forestry Research) for field work and CARN (Conservation Action Research Network) for complementary field work.

Acknowledgments: The authors thank Adeline Fayolle, Vivien Rossi and the LaBosystE of Bonaventure Sonké for their help in this research.
Conflicts of Interest: The authors declare no conflict of interest. The funders had no role in the design of the study; in the collection, analyses, or interpretation of data; in the writing of the manuscript, or in the decision to publish the results.

## References

1. Saatchi, S.S.; Marlier, M.; Chazdon, R.L.; Clark, D.B.; Russell, A.E. Impact of spatial variability of tropical forest structure on radar estimation of aboveground biomass. Remote Sens. Environ. 2011, 115, 2836-2849. [CrossRef]
2. Saatchi, S.S.; Harris, N.L.; Brown, S.; Lefsky, M.; Mitchard, E.T.A.; Salas, W.; Zutta, B.R.; Buermann, W.; Lewis, S.L.; Hagen, S.; et al. Benchmark map of forest carbon stocks in tropical regions across three continents. Proc. Natl. Acad. Sci. USA 2011, 108, 9899-9904. [CrossRef] [PubMed]
3. Asner, G.P.; Mascaro, J. Mapping tropical forest carbon: Calibrating plot estimates to a simple LiDAR metric. Remote Sens. Environ. 2014, 140, 614-624. [CrossRef]
4. Hubbell, S.P. The Unified Neutral Theory of Biodiversity and Biogeography; Princeton University Press: Princeton, NJ, USA, 2001; Volume 32, ISBN 0691021295.
5. Ricklefs, R.E. A comprehensive framework for global patterns in biodiversity. Ecol. Lett. 2004, 7, 1-15. [CrossRef]
6. Hardy, O.; Sonké, B. Spatial pattern analysis of tree species distribution in a tropical rain forest of Cameroon: Assessing the role of limited dispersal and niche differentiation. For. Ecol. Manag. 2004, 197, 191-202. [CrossRef]
7. Fine, P.V.A.; García-Villacorta, R.; Pitman, N.C.A.; Mesones, I.; Kembel, S.W. A Floristic Study of the White-Sand Forests of Peru. Ann. Mo. Bot. Gard. 2010, 97, 283-305. [CrossRef]
8. Anderson, A.B. White-sand vegetation of Brazilian Amazonia. Biotropica 1981,13,199-210. [CrossRef]
9. de Assis, R.L.; Wittmann, F. Forest structure and tree species composition of the understory of two central Amazonian várzea forests of contrasting flood heights. Flora Morphol. Distrib. Funct. Ecol. Plants 2011, 206, 251-260. [CrossRef]
10. Wittmann, F.; Householder, E.; Piedade, M.T.F.; De Assis, R.L.; Schöngart, J.; Parolin, P.; Junk, W.J. Habitat specifity, endemism and the neotropical distribution of Amazonian white-water floodplain trees. Ecography (Cop.) 2013, 36, 690-707. [CrossRef]
11. Fortunel, C.; Paine, C.E.T.; Fine, P.V.A.; Kraft, N.J.B.; Baraloto, C. Environmental factors predict community functional composition in Amazonian forests. J. Ecol. 2014, 102, 145-155. [CrossRef]
12. Watt, A.S. Pattern and Process in the Plant Community. J. Ecol. 1947, 35, 1-22. [CrossRef]
13. Réjou-Méchain, M.; Flores, O.; Pélissier, R.; Fayolle, A.; Fauvet, N.; Gourlet-Fleury, S. Tropical tree assembly depends on the interactions between successional and soil filtering processes. Glob. Ecol. Biogeogr. 2014, 23, 1440-1449. [CrossRef]
14. Connell, J.H.; Lowman, M.D. Low-diversity tropical rain forests: Some possible mechanisms for their existence. Am. Nat. 1989, 134, 88-119. [CrossRef]
15. Hart, T.B. Monospecific dominance in tropical rain forests. Trends Ecol. Evol. 1990, 5, 6-11. [CrossRef]
16. ter Steege, H.; Henkel, T.W.; Helal, N.; Marimon, B.S.; Marimon-Junior, B.H.; Huth, A.; Groeneveld, J.; Sabatier, D.; de Coelho, L.S.; de Filho, D.A.L.; et al. Rarity of monodominance in hyperdiverse Amazonian forests. Sci. Rep. 2019, 9, 1-15. [CrossRef]
17. Peh, K.S.H.; Lewis, S.L.; Lloyd, J. Mechanisms of monodominance in diverse tropical tree-dominated systems. J. Ecol. 2011, 99, 891-898. [CrossRef]
18. Janzen, D.H. Herbivores and the number of tree species in tropical forests. Am. Nat. 1970, 104, 501-528. [CrossRef]
19. Connell, J. On the role of natural enemies in preventing competitive exclusion in some marine animals and in rain forest trees. In Dynamics of Populations; den Boer, P., Gradwell, G., Eds.; Dynamics of Populations: Oosterbeek, The Netherlands, 1971; Volume 298, pp. 298-312.
20. Hart, T.; Hart, J.; Murphy, P. Monodominant and Species-Rich Forests of the Humid Tropics: Causes for Their Co-Occurrence. Am. Nat. 1989, 133, 613-633. [CrossRef]
21. Torti, S.D.; Coley, P.D.; Kursar, T.A. Causes and consequences of monodominance in tropical lowland forests. Am. Nat. 2001, 157, 141-153. [CrossRef] [PubMed]
22. Richards, P. The Tropical Rain Forest: An Ecological Study, 2nd ed.; Cambridge University Press: Cambridge, UK, 1996; ISBN 0-521-42194-2.
23. Martijena, N.E. Soil properties and seedling establishment in soils from monodominant and high-diversity stands of the tropical deciduous forests of Mexico. J. Biogeogr. 1998, 25, 707-719. [CrossRef]
24. Peh, K.S.H.; Sonké, B.; Lloyd, J.; Quesada, C.A.; Lewis, S.L. Soil does not explain monodominance in a Central African tropical forest. PLoS ONE 2011, 6, e16996. [CrossRef]
25. Lokonda, M.; Freycon, V.; Gourlet-Fleury, S.; Kombele, F. Are soils under monodominant Gilbertiodendron dewevrei and under adjacent mixed forests similar? A case study in the Democratic Republic of Congo. J. Trop. Ecol. 2018, 34, 176-185. [CrossRef]
26. Barbier, N.; Libalah, M.B.; Katembo, J.; Ploton, P.; Droissart, V.; Texier, N.; Kamdem, N.G.; Viennois, G.; Pélissier, R.; Couteron, P.; et al. Pistes pour l'étude de la distribution des peuplements de Gilbertiodendron dewevrei comme signature des impacts climatiques ou anthropiques anciens. In Pour Une écologie Historique en Afrique Centrale; AUF-IRD: Bondy, France, 2017; pp. 157-170.
27. Letouzey, R. Quelques exemples camerounais de liaison possible entre phénomènes géologiques et végétation. Bothalia 1983, 14, 739-744. [CrossRef]
28. Vleminckx, J.; Drouet, T.; Amani, C.; Lisingo, J.; Lejoly, J.;Hardy, O.J. Impact of fine-scale edaphic heterogeneity on tree species assembly in a central African rainforest. J. Veg. Sci. 2015, 26, 134-144. [CrossRef]
29. Libalah, M.B.; Droissart, V.; Sonké, B.; Barbier, N.; Dauby, G.; Fortunel, C.; Kamdem, G.; Kamdem, N.; Lewis, S.L.; Mofack, G.I.; et al. Additive influences of soil and climate gradients drive tree community composition of Central African rainforests. J. Veg. Sci. 2020, under review.
30. Umunay, P.M.; Gregoire, T.G.; Ashton, M.S. Estimating biomass and carbon for Gilbertiodendron dewevrei (De Wild) Leonard, a dominant canopy tree of African tropical Rainforest: Implications for policies on carbon sequestration. For. Ecol. Manag. 2017, 404, 31-44. [CrossRef]
31. Djuikouo, M.N.K.; Doucet, J.L.; Nguembou, C.K.; Lewis, S.L.; Sonké, B. Diversity and aboveground biomass in three tropical forest types in the Dja Biosphere Reserve, Cameroon. Afr. J. Ecol. 2010, 48, 1053-1063. [CrossRef]
32. Lewis, S.L.; Sonke, B.; Sunderland, T.; Begne, S.K.; Lopez-Gonzalez, G.; van der Heijden, G.M.F.; Phillips, O.L.; Affum-Baffoe, K.; Baker, T.R.; Banin, L.; et al. Above-ground biomass and structure of 260 African tropical forests. Philos. Trans. R. Soc. B Biol. Sci. 2013, 368. [CrossRef]
33. Makana, J.-R.; Hart, T.B.; Hibbs, D.E.; Condit, R.S. Stand structure and species diversity in the Ituri forest dynamics plots: A comparison of monodominant and mixed forest stands. In Tropical Forest Diversity and Dynamism; Losos, E.C., Leigh, E.C., Eds.; University of Chicago Press: Chicago, IL, USA, 2004; Volume 30, pp. 159-174. ISBN 0226493466.
34. Köppen, W. Versuch einer Klassifikation der Klimate, vorzugsweise nach ihren Beziehungen zur Pflanzenwelt. Geogr. Z. 1900, 11, 593-611.
35. Kottek, M.; Grieser, J.; Beck, C.; Rudolf, B.; Rubel, F. World Map of the Köppen-Geiger climate classification updated. Meteorol. Z. 2006, 15, 259-263. [CrossRef]
36. Fick, S.E.; Hijmans, R.J. WorldClim 2: New 1-km spatial resolution climate surfaces for global land areas. Int. J. Climatol. 2017, 37, 4302-4315. [CrossRef]
37. Hijmans, R.J.; Cameron, S.E.; Parra, J.L.; Jones, P.G.; Jarvis, A. Very high resolution interpolated climate surfaces for global land areas. Int. J. Climatol. 2005, 25, 1965-1978. [CrossRef]
38. Réjou-Méchain, M.; Barbier, N.; Couteron, P.; Ploton, P.; Vincent, G.; Herold, M.; Mermoz, S.; Saatchi, S.; Chave, J.; de Boissieu, F.; et al. Upscaling Forest Biomass from Field to Satellite Measurements: Sources of Errors and Ways to Reduce Them. Surv. Geophys. 2019, 40, 881-911. [CrossRef]
39. Larjavaara, M.; Muller-Landau, H.C. Measuring tree height: A quantitative comparison of two common field methods in a moist tropical forest. Methods Ecol. Evol. 2013, 4, 793-801. [CrossRef]
40. APG IV. An update of the Angiosperm Phylogeny Group classification for the orders and families of flowering plants: APG IV. Bot. J. Linn. Soc. 2016, 181, 1-20.
41. African Plants Database (version 3.4.0) Conservatoire et Jardin botaniques de la VIlle de Genève and South African National Biodiversity Institute Pretoria. Available online: https://www.ville-ge.ch/musinfo/bd/cjb/ africa/recherche.php (accessed on 20 January 2017).
42. Rennó, C.D.; Nobre, A.D.; Cuartas, L.A.; Soares, J.V.; Hodnett, M.G.; Tomasella, J.; Waterloo, M.J. HAND, a new terrain descriptor using SRTM-DEM: Mapping terra-firme rainforest environments in Amazonia. Remote Sens. Environ. 2008, 112, 3469-3481. [CrossRef]
43. Quinn, P.F.; Beven, K.J.; Lamb, R. The in $(a / \tan / \beta)$ index: How to calculate it and how to use it within the topmodel framework. Hydrol. Process. 1995, 9, 161-182. [CrossRef]
44. Böhner, J.; Selige, T. Spatial prediction of soil attributes using terrain analysis and climate regionalisation. SAGA-Anal. Model. Appl. Göttinger Aeographische Abhandlungen 2006, 115, 13-28.
45. Feldpausch, T.R.; Lloyd, J.; Lewis, S.L.; Brienen, R.J.W.; Gloor, M.; Monteagudo Mendoza, A.; Lopez-Gonzalez, G.; Banin, L.; Abu Salim, K.; Affum-Baffoe, K.; et al. Tree height integrated into pantropical forest biomass estimates. Biogeosciences 2012, 9, 3381-3403. [CrossRef]
46. Réjou-Méchain, M.; Tanguy, A.; Piponiot, C.; Chave, J.; Hérault, B. Biomass: An R Package for Estimating Above-Ground Biomass and Its Uncertainty in Tropical Forests. Methods Ecol. Evol. 2017, 8, 1163-1167. [CrossRef]
47. Chave, J.; Réjou-Méchain, M.; Búrquez, A.; Chidumayo, E.; Colgan, M.S.; Delitti, W.B.C.; Duque, A.; Eid, T.; Fearnside, P.M.; Goodman, R.C.; et al. Improved allometric models to estimate the aboveground biomass of tropical trees. Glob. Chang. Biol. 2014, 20, 3177-3190. [CrossRef]
48. Zanne, A.E.; Lopez-Gonzalez, G.; Coomes, D.A.A.; Ilic, J.; Jansen, S.; Lewis, S.L.S.L.; Miller, R.B.B.; Swenson, N.G.G.; Wiemann, M.C.C.; Chave, J. Global wood density database. Dryad 2009, 235, 33.
49. Pélissier, R.; Couteron, P.; Dray, S.; Sabatier, D. Consistency between ordination techniques and diversity measurements: Two strategies for species occurrence data. Ecology 2003, 84, 242-251. [CrossRef]
50. Gimaret-Carpentier, C.; Chessel, D.; Pascal, J.P. Non-symmetric correspondence analysis: An alternative for species occurrences data. Plant Ecol. 1998, 138, 97-112. [CrossRef]
51. Kroonenberg, P.M.; Lombardo, R. Nonsymmetric correspondence analysis: A tool for analysing contingency tables with a dependence structure. Multivar. Behav. Res. 1999, 34, 367-396. [CrossRef]
52. Oksanen, J.; Blanchet, F.; Kindt, R.; Legendre, P.; O'Hara, R. Vegan: Community Ecology Package; R Packag. Version 2.3-3; CRAN R-project: 2016. Available online: https://cran.r-project.org/web/packa (accessed on 22 December 2016).
53. Chase, J.M.; McGill, B.J.; McGlinn, D.J.; May, F.; Blowes, S.A.; Xiao, X.; Knight, T.M.; Purschke, O.; Gotelli, N.J. Embracing scale-dependence to achieve a deeper understanding of biodiversity and its change across communities. Ecol. Lett. 2018, 21, 1737-1751. [CrossRef] [PubMed]
54. McGlinn, D.J.; Xiao, X.; May, F.; Gotelli, N.J.; Engel, T.; Blowes, S.A.; Knight, T.M.; Purschke, O.; Chase, J.M.; McGill, B.J. Measurement of Biodiversity (MoB): A method to separate the scale-dependent effects of species abundance distribution, density, and aggregation on diversity change. Methods Ecol. Evol. 2019, 10, 258-269. [CrossRef]
55. Dauby, G.; Hardy, O.J. Sampled-based estimation of diversity sensu stricto by transforming Hurlbert diversities into effective number of species. Ecography (Cop.) 2012, 35, 661-672. [CrossRef]
56. Simpson, E.H. Measurement of diversity. Nature 1949, 163, 688. [CrossRef]
57. Jost, L. Entropy and diversity. Oikos 2006, 113, 363-375. [CrossRef]
58. O'Brien, S.T.; Hubbell, S.P.; Spiro, P.; Condit, R.; Foster, R.B. Diameter, height, crown, and age relationships in eight neotropical tree species. Ecology 1995, 76, 1926-1939. [CrossRef]
59. Newbery, D.M.; Alexander, I.J.; Rother, J.A. Phosphorus dynamics in a lowland African rain forest: The influence of ectomycorrhizal trees. Ecol. Monogr. 1997, 67, 367-409. [CrossRef]
60. Hart, T.B. The Ecology of a Single-Species-Dominant Forest and a Mixed Forest in Zaire. Unpublished Ph.D. Thesis, Michigan State University, East Lansing, MI, USA, 1985.
61. Fayolle, A.; Loubota Panzou, G.J.; Drouet, T.; Swaine, M.D.; Bauwens, S.; Vleminckx, J.; Biwole, A.; Lejeune, P.; Doucet, J.-L. Taller trees, denser stands and greater biomass in semi-deciduous than in evergreen lowland central African forests. For. Ecol. Manag. 2016, 374, 42-50. [CrossRef]
62. Chave, J.; Andalo, C.; Brown, S.; Cairns, M.A.; Chambers, J.Q.; Eamus, D.; Fölster, H.; Fromard, F.; Higuchi, N.; Kira, T.; et al. Tree allometry and improved estimation of carbon stocks and balance in tropical forests. Oecologia 2005, 145, 87-99. [CrossRef]
63. Brown, S. Estimating Biomass and Biomass Change of Tropical Forests: A Primer; FAO Forest; Food \& Agriculture Org.: Rome, Italy, 1997; ISBN 9251039550.
64. Wilson, J.B.; Agnew, A.D.Q. Positive-feedback Switches in Plant Communities. Adv. Ecol. Res. 1992, 23, 263-336.
65. Henkel, T.W. Monodominance in the ectomycorrhizal Dicymbe corymbosa (Caesalpiniaceae) from Guyana. J. Trop. Ecol. 2003, 19, 417-437. [CrossRef]
66. Torti, S.D.; Coley, P.D. Tropical Monodominance: A Preliminary Test of the Ectomycorrhizal Hypothesis Biotropica 1999, 31, 220-228. [CrossRef]
67. Green, J.J.; Newbery, D.M. Light and seed size affect establishment of grove-forming ectomycorrhizal rain forest tree species. New Phytol. 2001, 151, 271-289. [CrossRef]
68. Letouzey, R. Carte Phytogéographique du Cameroun, 1:500 000, 8 Feuillets +5 Notices; Institut de la Carte Internationale de la Végétation: Toulouse, France, 1985.
© 2020 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (http://creativecommons.org/licenses/by/4.0/).
