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International Variability of Ages at Menarche and Menopause: Patterns and Main Determinants

FRÉDÉRIC THOMAS,¹ FRANÇOIS RENAUD,¹ ERIC BENEFICE,² THIERRY DE MEEÛS,¹ AND JEAN-FRANÇOIS GUEGAN¹

Abstract The purpose of this study was to review published studies on the variability of age at menarche and age at menopause throughout the world, and to identify the main causes for age variation in the timing of these events. We first present a summary table including mean (or median) values of the age at menarche in 67 countries, and of the age at menopause in 26 countries. General linear models showed that mean age at menarche was strongly linked to the mean female life expectancy, suggesting that one or several variables responsible for inequalities in longevity similarly influenced the onset of menarche. A closer examination of the data revealed that among several variables reflecting living conditions, the factors best explaining the variation in age at menarche were adult illiteracy rate and vegetable calorie consumption. Because adult illiteracy rate has some correlation with the age at which children are involved in physical activities that can be detrimental in terms of energy expenditure, our results suggest that age at menarche reflects more a trend in energy balance than merely nutritional status. In addition, we found the main determinant of age at menopause to be the mean fertility. This study thus suggests that, on a large scale, age at menarche is mainly determined by extrinsic factors such as living conditions, while age at menopause seems to be mainly influenced by intrinsic factors such as the reproductive history of individuals. Finally, these findings suggest that human patterns cannot be addressed solely by traditional, small-scale investigations on single populations. Rather, complementary research on a larger scale, such as this study, may be more appropriate in defining some interesting applications to the practical problems of human ecology.

Menarche (first menstrual period) and menopause (end of menstruation) are the two major components in the reproductive life of women, since the interval between the two events determines the natural reproductive period during which females can procreate. Because these two biological traits have important cultural, social, and epidemiological implications, increasing attention has been recently

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devoted by scientists to understanding the causes of age variations in the timing of these events. Although results are not always consistent from one study to another, several factors have been shown to significantly influence age at menarche and at menopause, such as genetic parameters (Danker-Hopfe and Delibalta 1990; Kaprio et al. 1995; Treolar et al. 1998), socioeconomic conditions (Belmaker 1982; Luoto et al. 1994), general health and life-style (Parazzini et al. 1992; Brown et al. 1996), nutritional status (Osteria 1983; Riley 1994; Simondon et al. 1997), seasonality (Boldsen 1992), physical activity (Malina 1983; Baker 1985), and altitude level (Beall 1983; Kapoor and Kapoor 1986; Gonzales and Villena 1996).

Mean ages at menarche and at menopause vary substantially between women across different countries or across different ethnic groups (Belitz 1977; Gray and Doyle 1983; Hunt and Newcomer 1984; Danker 1986; Ulijaszek et al. 1991; Flint 1997; Morabia et al. 1998). Reasons behind this international variability remain poorly understood, mainly because few comparative analyses have been conducted on such a large scale (but see Morabia et al. 1998). Most investigations of the relationships between ages at menarche and menopause and their causes have been studied on a small scale by social scientists, anthropologists, or public health epidemiologists. However, the variables affecting the timing of these events within populations do not necessarily explain differences between populations. In order to characterize patterns in the variation of ages at both menarche and menopause across populations, it is necessary to conduct examinations on large spatial and/or temporal scales (i.e., secular trends) in order to place the findings within a broader perspective. This approach is possible by adopting the techniques of "macroecology," which try to rise above the many details of local patterns to find a larger picture from which a kind of statistical order emerges (Lawton 1999).

In this study, we collected data on age at menarche and age at menopause in different countries from studies published worldwide. In order to understand the determinants of the timing of these events on this large spatial scale, we then attempted to regress the variability observed within these traits to various geographical, socioeconomic, cultural, and biological variables.

Material and Methods

Data. A literature search was performed for publications on age at menarche and age at menopause. All papers that displayed a measure (mean or median) for a given country of the timings of these events were retained for the analysis. When several values were given for the same country, we used the mean of these values. Some published data refer to a limited part of a population (e.g., a given city), and an important assumption we made is that they are representative of the country (but see the Discussion section). Only when conditions for parametric statistical analyses were violated (homogeneity of variances, error distribution), did we use log-transformed values (Zar 1996). The normality of error distribution

was tested using Shapiro-Wilks statistics. Thus, age at menarche, mean fertility, and illiteracy rate values were log-transformed while age at menopause values were kept unchanged.

Socioeconomic and geographical parameters were obtained from the 1992 world population data sheet (Jones 1990) and from World Atlas v. 2.1.0©: (1) the female life expectancy at birth (in years), (2) the mean fertility (number of offspring born to a woman throughout the child-bearing years), (3) the per capita gross national product (GNP in US\$ a year), (4) the population density (number of people per square kilometer), (5) the animal and vegetable calorie consumption per person and per day, (6) the mean latitude in degree and minutes, which refers to the value taken at the geographical center of each country, (7) the hemisphere of residency (northern is coded 1 and southern is coded 2). Data on mean female age at marriage were available from the Women's Indicators and Statistics data base (Wistat, version 3, United Nations publication). Adult illiteracy rates (expressed as a percentage) for each country were obtained from the United Nations Educational, Scientific and Cultural Organization.

Statistical Analyses. The possible sources of environmental heterogeneity at the largest spatial scale are extraordinarily various and thus cannot be simultaneously introduced in the models when attempting to explain the variation in ages at menarche and menopause. Thus, the contribution of different potential explanatory variables and their interaction terms to the dependent variables was first derived by general linear modeling (GLM) regressions with all independent variables retained in the final models (Zar 1996). In addition, minimal models were also selected after backward (and forward) stepwise elimination protocol. At each iteration, the variable showing the lowest partial regression coefficient with the dependent variable was removed from the model if its relationship was not significant at the 5% level (Zar 1996). When no variable could be removed from the model, that is to say when all variables are significant, the procedure was deemed finished. We used the tolerance option with a value of 0.05, which protects against constructing highly multicollinear models (Venables and Ripley 1994).

Age at Menarche. Since geographical factors (and/or their correlates) might influence the variation of age at menarche, we introduced in GLM regressions the mean latitude and the hemisphere of residency. To deal with the possible influence of demographic and economic variables, we also considered in the GLM regressions the population density and the per capita gross national product. Given the two kinds of values obtained for average age at menarche (i.e., mean or median), we controlled for this potential bias by introducing a dichotomous variable (named sampling bias, coded 1 for mean values and 2 for median values). The nonsignificance of this variable would indicate that we can pool mean and median values together, analyzing the entire data set simultaneously despite the heterogeneity of published data reporting average values. Finally, we introduced in the analysis the female life expectancy at birth, assuming that other living condi-

Table 1. Mean (or Median*) Age at Menarche and at Menopause in the Different Available Countries

| <i>Country</i> | <i>Age at Menarche (in Years)</i> | <i>Reference</i> | <i>Age at Menopause (in Years)</i> | <i>Reference</i> |
|----------------|-----------------------------------|--------------------------------------|------------------------------------|------------------------------|
| Algeria | 14.3 | Grassivaro-Gallo and Florio (1993) | — | — |
| Argentina | 12.59 | Zurlo de Mirotti et al. (1995) | — | — |
| Australia | 13.0* | Morabia et al. (1998) | 50.4 | Walsh (1978) |
| Bangladesh | 15.8 | Riley and Khan (1993) | — | — |
| Belgium | 13.0* | Vercauteren and Susanne (1984) | — | — |
| Britain | 13.3 | Mascie-Taylor and Boldsen (1986) | — | — |
| Cameroon | 14.61 | Biyong et al. (1985) | — | — |
| Canary Islands | — | — | 48.6 | Sosa Henriquez et al. (1994) |
| Chile | 13.0* | Morabia et al. (1998) | 50.0* | Morabia et al. (1998) |
| China | 12.38* | Huen et al. (1997) | 49.0* | Morabia et al. (1998) |
| Colombia | 12.8 | Pardo and Uriza (1991) | 50.0* | Morabia et al. (1998) |
| Congo-Brazza | 12.0 | Samba (1982) | — | — |
| Congo-Kinshasa | 13.83 | Rashid-Tozin et al. (1984) | — | — |
| Cuba | 13.01 | Jordan-Rodriguez et al. (1980) | — | — |
| Czechoslovakia | 14.6 | Magursky et al. (1975) | 51.2 | Magursky et al. (1975) |
| Denmark | 13.0 | Helm and Grolund (1998) | — | — |
| Dominican Rep. | 12.6 | Mancebo (1990) | — | — |
| Egypt | 13.2 | Attallah (1978) | — | — |
| Finland | 13.2 | Dahlstrom et al. (1984) | 51.0* | Luoto et al. (1994) |
| France | 13.05* | Crognier and Tavares Da Rocha (1979) | 52.0 | Salat-Baroux (1980) |

| | | | | |
|-----------------------|--------|--|-------|---------------------------------|
| East Germany | 14.0* | Morabia et al. (1998) | — | — |
| Ghana | 13.98 | Adadevoh et al. (1989) | 48.05 | Kwawukume et al. (1993) |
| Greece | 12.0 | Pentzos-Daponte and Grefen-Peters (1984) | — | — |
| Guatemala | 13.75 | Khan et al. (1995) | — | — |
| Haiti | 15.37 | Barnes-Josiah and Augustin (1995) | — | — |
| Hungary | 12.9 | Dober and Kiralyfalvi (1993) | — | — |
| Iceland | 13.06 | Macgusson (1978) | — | — |
| India (Punjab) | 14.31 | Singh and Ahuja (1980) | 44.6 | Singh and Ahuja (1980) |
| Indonesia | 13.0 | Samsudin (1990) | 50.5 | Samil and Wishnuwardhani (1994) |
| Ireland | 13.52 | Hoey et al. (1986) | — | — |
| Israel | 13.29 | Belmaker (1982) | — | — |
| Italy | 12.2 | Zoppi (1992) | — | — |
| Jamaica | 13.1 | Jamaica National Family Planning (1988) | — | — |
| Japan | 12.5 | Nakamura et al. (1986) | 49.3 | Kono et al. (1990) |
| Kenya | 14.4 | Rogo et al. (1987) | — | — |
| Malaysia | 14.2 | Wilson (1985) | 50.7 | Ismael (1994) |
| Mexico | 12.4 | Garcia-Baltazar et al. (1993) | 46.5 | Garrido-Latorre et al. (1996) |
| Morocco | 13.75* | Loukid et al. (1996) | — | — |
| Nepal (high altitude) | 16.2* | Beall (1983) | 46.8* | Beall (1983) |
| New-Zealand | 12.9 | St George et al. (1994) | — | — |
| Nicaragua | 14.0 | Guido et al. (1971) | — | — |
| Nigeria | 15.0* | Morabia et al. (1998) | 48.4 | Okonofua et al. (1990) |
| Norway | 13.2* | Nafstad et al. (1995) | — | — |
| Papua New Guinea | 15.8 | Groos and Smith (1992) | — | — |
| Peru | 13.23 | Soto-Caceres and Guevara-Servigon (1988) | — | — |
| Philippines | 13.6 | Zablan (1988) | 48.0 | Ramoso-Jalbuena (1994) |

Table 1. Continued

| <i>Country</i> | <i>Age at Menarche (in Years)</i> | <i>Reference</i> | <i>Age at Menopause (in Years)</i> | <i>Reference</i> |
|----------------------------|-----------------------------------|---|------------------------------------|--------------------------|
| Poland | 13.06 | Laska-Mierzejewska et al. (1982) | — | — |
| Roumania | 13.47 | Stukovsky et al. (1967) | — | — |
| Russia | 13.0 | Iampol'skaia (1997) | 49.0 | Balan (1995) |
| Sardinia | 12.78* | Floris et al. (1987) | — | — |
| Senegal | 16.1 | Simondon et al. (1997) | — | — |
| Somalia | 14.78 | Gallo (1975) | — | — |
| South Africa (black women) | — | — | 49.2 | Walker et al. (1984) |
| Southern Korea | 13.9 | Kim et al. (1986) | — | — |
| Spain | 12.31 | de la Puente et al. (1997) | — | — |
| Sri Lanka | 13.5 | Balasuriya and Fernando (1983) | — | — |
| Sudan | 13.75 | Attallah et al. (1983) | — | — |
| Sweden | 13.09 | Furu (1976) | 50.9* | Hagstad (1988) |
| Switzerland | 13.0 | Morabia et al. (1996) | 50.0 | Morabia et al. (1996) |
| Tahiti | 12.75* | Ducros and Ducros (1987) | — | — |
| Taiwan | 13.6 | Chow et al. (1997) | 49.5 | Chow et al. (1997) |
| Tanzania | 15.21 | Hautvast (1971) | — | — |
| Thailand | 12.3 | Piya-Anant et al. (1997) | 50.3 | Tunghaisal et al. (1991) |
| Turkey | 13.28 | Vicdan et al. (1996) | 47.8 | Carda et al. (1998) |
| United Arab Emirates | — | — | 47.3 | Rizk et al. (1998) |
| USA | 12.8 | Malina and Bouchard (1991) | 51.3* | Kato et al. (1998) |
| Venezuela | 12.68* | Farid-Coupal et al. (1981) | — | — |
| Yemen | 14.4 | Yemen Arab Republic Fertility Survey (1979) | — | — |
| Zambia | 13.7 | Katzarski et al. (1980) | — | — |
| Zimbabwe | 13.5 | Mbizvo et al. (1995) | — | — |

tions able to influence the onset of menarche should also be reflected in life expectancy.

Age at Menopause. In order to assess the predictors of age at menopause, we considered two additional independent variables. Because age at menopause could be influenced by the length of active sexual life, we introduced in the GLM regressions the mean female age at marriage for each country, which roughly determines in most societies the beginning of reproductive life for females. Then, because there could be a trade-off between reproductive effort and length of sexual life (e.g., Westendorp and Kirkwood 1998; Thomas et al. 2000), we also introduced the mean fertility of each country. As before, we again considered the potential artefact exerted by the heterogeneity of published data in reporting age at menopause values (i.e., mean coded 1, and median coded 2).

Life Expectancy. When female life expectancy had a significant influence in the two previous analyses, we performed further analyses aimed at determining the possible causes behind this relationship. Among the numerous factors able to influence life expectancies, we considered those having a priori been identified potentially to affect the timing of reproductive events. Among these factors, we focused on two nutritional variables (animal and vegetable calorie consumption per person and per day) and adult illiteracy rate. Illiteracy rate is here used as an estimator of the mean level of energy expenditure resulting from physical activity. Indeed, adult illiteracy rate is usually positively correlated with child labor (Parker 1997; Psacharopoulos 1997) and with the nature of professional activities subsequently exerted by individuals (Rougerie and Courtois 1997). We first attempted to explain the variability in life expectancies by these variables. Following this analysis, variables relevant to explain life expectancy variations were used as predictors when attempting to explain the timing of reproductive events.

Analyses were performed using the S-Plus statistical package (MathSoft, Inc. 1999; Venables and Ripley 1994).

Results

Data on age at menarche and age at menopause were obtained for 67 and 26 countries, respectively (Table 1). Overall, the mean age of menarche is 13.53 years ($SD \pm 0.98$) and the mean age of menopause is 49.24 years ($SD \pm 1.73$).

Values for the independent variables used to explain the international variation in age at menarche were obtained for a subset of 58 countries out of the 67. Age at menarche was only related to female life expectancy (Figure 1A) when all potential influencing variables were considered (Table 2, $N = 58$, multiple adj. $R^2 = 0.535$, $p < 0.001$). Both backward and forward elimination procedures yielded similar results, i.e., single effect of female life span on age at menarche ($N = 58$, simple adj. $r^2 = 0.420$, $p < 0.00001$).

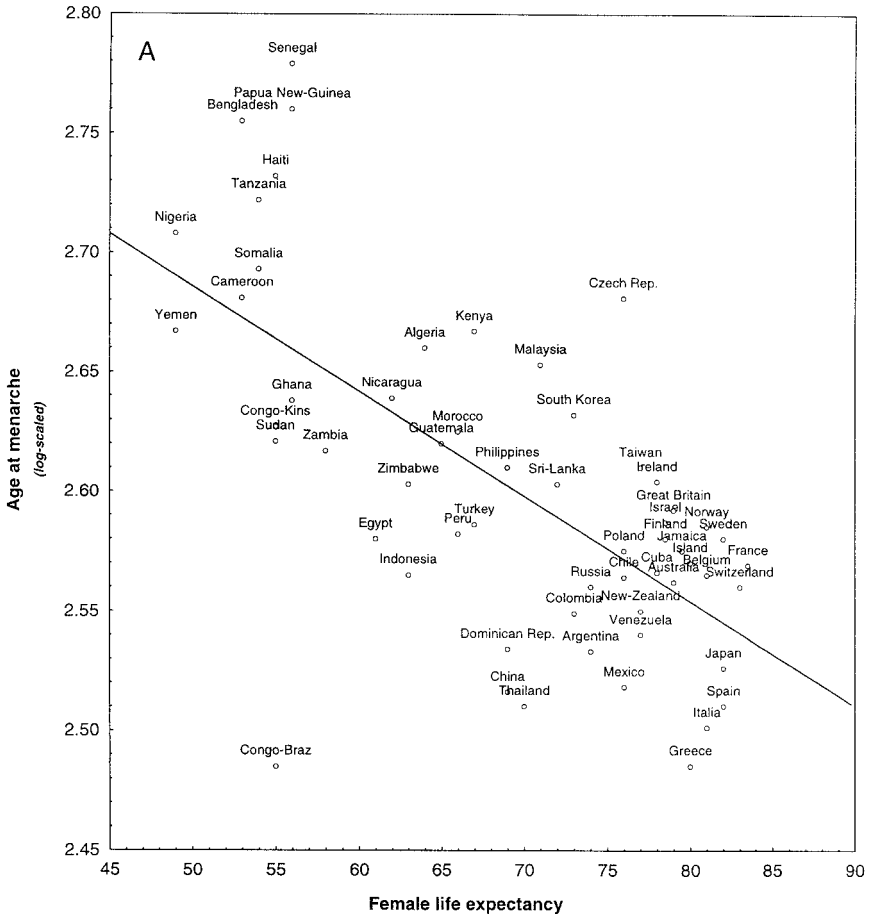


Figure 1. Relationships between age at menarche or age at menopause and their most relevant explanatory variables. **A**, Age at menarche and female life expectancy ($n = 58$, $r^2 = 0.434$, $y = -0.004x + 2.906$, $p < 0.00001$); **B**, Age at menarche and vegetable calorie consumption per person and per day ($n = 54$, $r^2 = 0.150$, $y = -0.002x + 2.781$, $p = 0.004$); **C**, Age at menarche and adult illiteracy rate ($n = 38$, $r^2 = 0.537$, $y = 0.033x + 2.539$, $p < 0.00001$); **D**, Age at menopause and fertility ($n = 23$, $r^2 = 0.442$, $y = -2.065x + 51.688$, $p = 0.0054$).

When performing a GLM procedure to analyze the variation in life expectancy across countries (total model not shown), we found that this trait was mainly explained by the adult illiteracy rate ($p < 0.00001$), GNP and vegetable calorie consumption being only marginally significant ($p = 0.072$ and $p = 0.076$, respectively). A backward stepwise procedure led to two equivalent minimal models in which illiteracy rate was highly significant in both cases, and vegetable

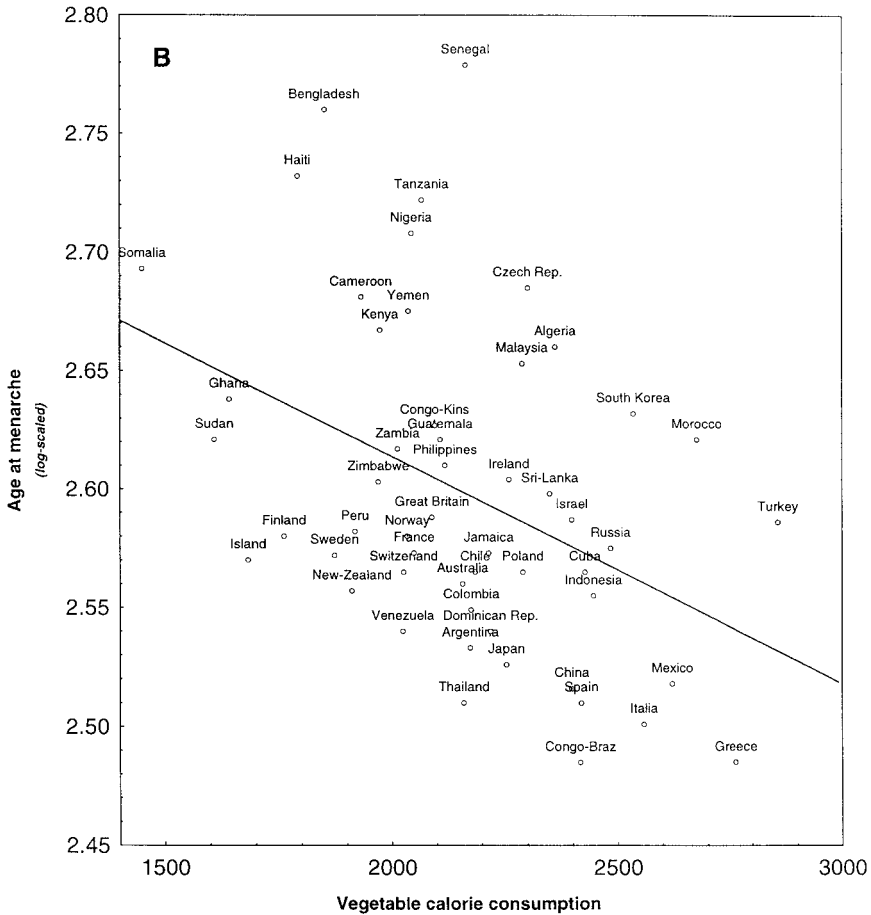


Figure 1. Continued

calorie consumption (Table 3, Model A, $N = 40$, multiple adj. $R^2 = 0.77$, $p < 0.0001$) or GNP (Table 3, Model B, $N = 41$, multiple adj. $R^2 = 0.79$, $p < 0.0001$) had an additional but exclusive contribution to the variance observed on female life expectancy.

Because GNP was already proposed as a possible predictor of age at menarche (see Table 2), we replaced the life expectancy variable by the two factors, i.e., illiteracy rate and vegetable calorie consumption, in the new explanatory model of age at menarche ($N = 37$, multiple adj. $R^2 = 0.57$, $p < 0.0001$). Both variables

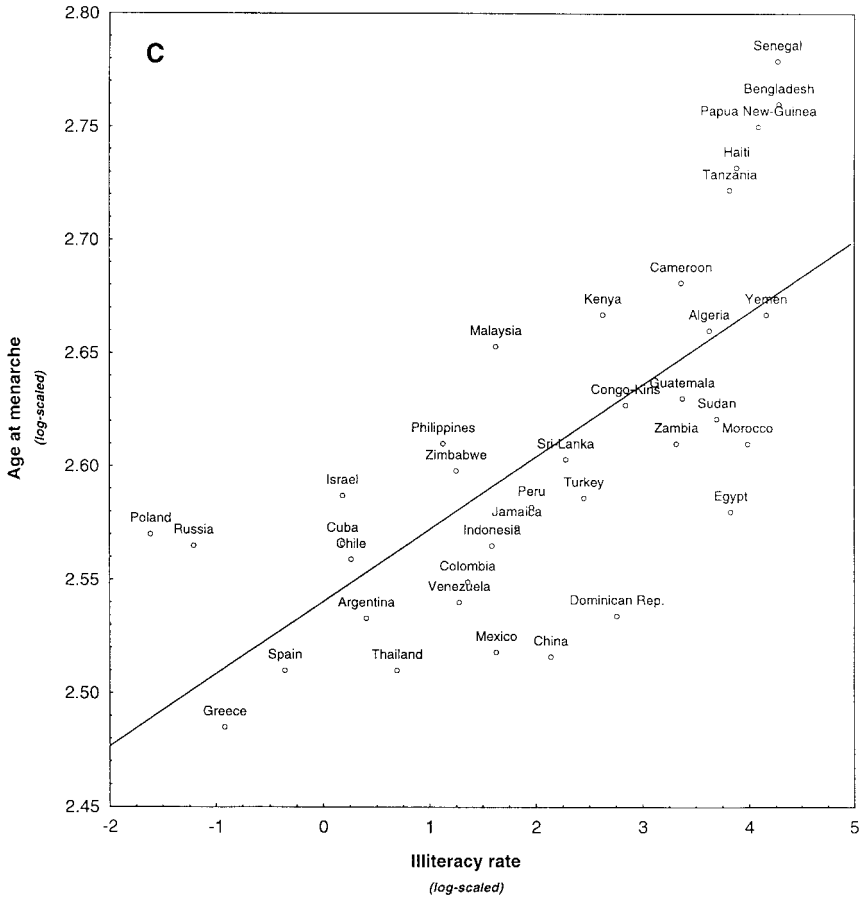


Figure 1. Continued

had a significant effect on age at menarche (Figures 1B and 1C; Table 4). Thus, at least two variables (i.e., vegetable calorie consumption and illiteracy rate) have the potential to influence living conditions at a level strong enough to modify both the onset of menarche and life expectancy.

Values for the independent variables used to explain the international variation in age at menopause were obtained for a subset of 21 countries. In a GLM model, only fertility had a significant effect (Table 5). Both backward and forward stepwise models yielded to similar results (Table 5; Figure 1D).

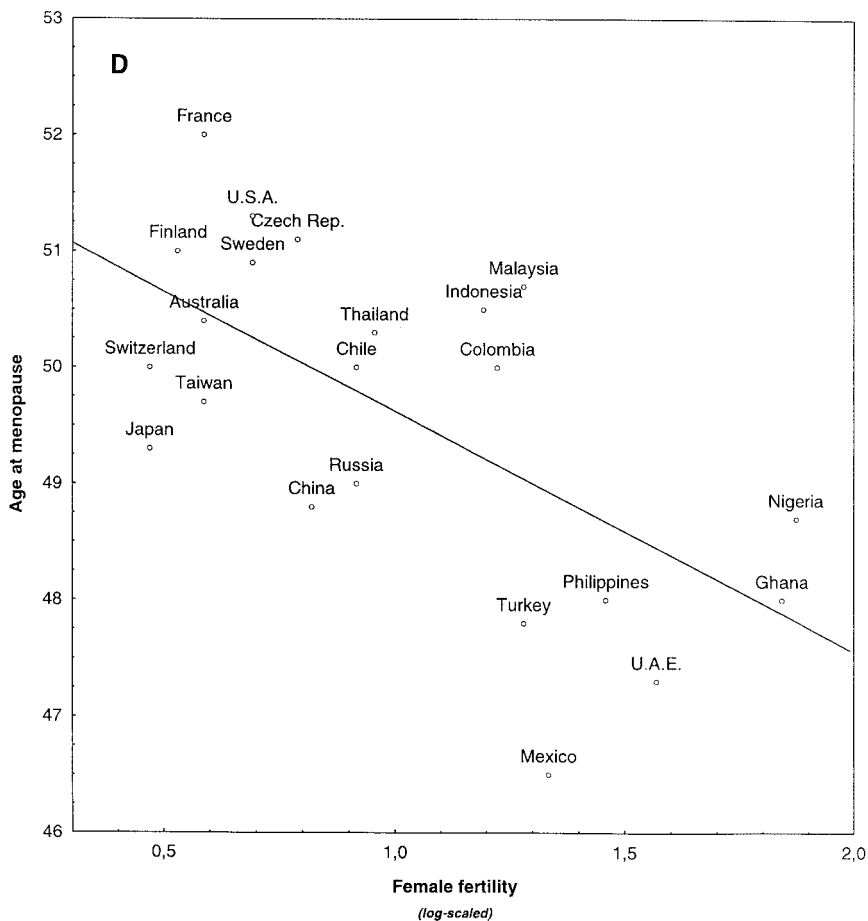


Figure 1. Continued

Discussion

It is frequently argued that comparative analyses using information from different sources may be inappropriate because data have been collected with different methods or they come from different sources. Although the argument is always applicable when no significant result is detected (i.e., data are not precise enough to detect a potentially significant result), it is unlikely to be relevant when significant trends are found, since a biological tendency has a priori no reason to be correlated to a background noise in the data set (Brown 1995; Rosenzweig 1995; Møller 1997; Lawton 1999).

Our results indicate that on a global scale, the age at menarche appears ear-

Table 2. General Linear Modelling for Age at Menarches (Log-Transformed) in Females across 58 Countries

| | <i>Parameter Estimate</i> | <i>Standard Error</i> | <i>df</i> | <i>F</i> | <i>Pr (F)</i> |
|----------------------------|---------------------------|-----------------------|-----------|----------|-------------------|
| Intercept | 1.415 | 0.195 | 1 | | |
| (1) Latitude | 0.071 | 0.050 | 1 | 4.015 | 0.0525 |
| (2) Female life expectancy | -0.008 | 0.005 | 1 | 54.031 | <0.0001 |
| (3) Population density | 0.002 | 0.030 | 1 | 2.462 | 0.1251 |
| (4) GNP | -0.013 | 0.034 | 1 | 0.063 | 0.8035 |
| (5) Hemisphere | 0.040 | 0.056 | 1 | 0.827 | 0.3690 |
| (6) Sampling bias | -0.008 | 0.042 | 1 | 0.022 | 0.8826 |
| (1):(2) | -0.001 | 0.001 | 1 | 1.850 | 0.1820 |
| (1):(3) | 0.004 | 0.007 | 1 | 0.253 | 0.6180 |
| (1):(4) | 0.005 | 0.005 | 1 | 6.672 | 0.0139 |
| (1):(5) | 0.019 | 0.009 | 1 | 0.998 | 0.3242 |
| (1):(6) | 0.008 | 0.007 | 1 | 1.485 | 0.2307 |
| (2):(3) | 0.001 | 0.001 | 1 | 0.087 | 0.7697 |
| (2):(4) | 0.001 | 0.001 | 1 | 2.257 | 0.1415 |
| (2):(5) | -0.003 | 0.002 | 1 | 1.235 | 0.2736 |
| (2):(6) | 0.001 | 0.001 | 1 | 0.100 | 0.7539 |
| (3):(4) | -0.005 | 0.006 | 1 | 2.057 | 0.1599 |
| (3):(5) | 0.010 | 0.010 | 1 | 0.404 | 0.5291 |
| (3):(6) | -0.002 | 0.004 | 1 | 0.990 | 0.3262 |
| (4):(5) | 0.009 | 0.011 | 1 | 0.168 | 0.6844 |
| (4):(6) | -0.003 | 0.005 | 1 | 0.590 | 0.4472 |
| (5):(6) | 0.010 | 0.010 | 1 | 1.176 | 0.2851 |

Note: Also given are the parameter estimate in regression, the standard error, the degree of freedom in analysis (*df*), the two-tailed probability value [*Pr (F)*] associated with the partial *F* statistics. The variable number 6, i.e. sampling bias, means that values of age at menarches have been categorized into mean values (coded 1) or median values (coded 2) in order to check for the influence of this sampling artefact on calculation. All two-way interaction terms of the expression (I):(J) have been introduced into the GLM procedure. Statistics in bold characters correspond to variable(s) retained in minimal models.

lier in countries where individuals experience the longest life expectancies. This result strongly suggests that the beginning of the capacity to reproduce in women is closely related to both biologically determined variables, i.e., genetic variability of life span, and culturally related variables, i.e., health and socioeconomic conditions. Previous studies have suggested this idea to explain the general decline in the menarcheal age in developed countries, concomitantly with the improvement of the living conditions during the last century (Nakamura et al. 1986; Tryggvadottir et al. 1994; Rees 1995). A close examination of our data reveals that at least two parameters reflecting living conditions influence variations of menarcheal age. Vegetable calorie consumption per person and per day has a small, but significant, influence on age at menarche, supporting the idea that good

Table 3. Two Alternative Minimal Linear Models for Explaining the Variability of Life Expectancy in Females across 40 (Model A) and 41 (Model B) Countries, Respectively

Model A

| | <i>Parameter Estimate</i> | <i>Standard Error</i> | <i>df</i> | <i>F Value</i> | <i>Pr (F)</i> |
|-----------------------|---------------------------|-----------------------|-----------|----------------|---------------|
| Intercept | 73.004 | 12.690 | 1 | | |
| (1) Illiteracy rate | -9.835 | 3.928 | 1 | 116.269 | <0.0001 |
| (2) Vegetable calorie | 0.002 | 0.005 | 1 | 10.828 | 0.0022 |
| (1):(2) | 0.002 | 0.002 | 1 | 2.045 | 0.1613 |

Model B

| | <i>Parameter Estimate</i> | <i>Standard Error</i> | <i>df</i> | <i>F Value</i> | <i>Pr (F)</i> |
|---------------------|---------------------------|-----------------------|-----------|----------------|---------------|
| Intercept | 67.466 | 7.884 | 1 | | |
| (1) Illiteracy rate | -9.611 | 2.532 | 1 | 140.332 | <0.0001 |
| (2) GNP | 1.153 | 0.947 | 1 | 12.882 | 0.0009 |
| (1):(2) | 0.839 | 0.331 | 1 | 6.421 | 0.0656 |

Note: Abbreviations are as in Table 2. The expression (1):(2) represents for model A the interaction between the woman illiteracy rate and the consumption in vegetable calorie across countries, and for model B the interaction between the woman illiteracy rate and the global net incomes across countries.

nutritional conditions favor early menarche (Frisch 1982, 1985; Warren 1983; Haq 1984). Poor nutrition is known to alter the ratio of lean mass to body fat and to delay the onset of menarche (Baker 1985). The strongest factor remains, however, the illiteracy rate. Illiteracy rate is usually correlated with child labor (Parker 1997; Psacharopoulos 1997), which is itself frequently responsible for reduced growth and reduced development of children (Parker 1997). Indeed, child labor often results in physical activities that require a higher level of energy expenditure than that required at school. Excessive physical exercise is, for instance, known in athletic women to disrupt fat accumulation and, subsequently, to delay the onset of menarche (Baker 1985). Although further information would be necessary to address this point in detail, our results support the hypothesis that ovarian function reflects trends in energy balance, not merely nutritional status (Ellison 1990).

Age at menopause seems to be linked with variables other than those influencing life expectancy. The negative effect of the fertility variable on the age at menopause suggests the possible existence of reproductive costs in humans. For instance, early and high investments into reproduction may result in shorter longevity (Westendorp and Kirkwood 1998; Thomas et al. 2000). However, other studies have not found a significant relationship between fertility and age at

Table 4. General Linear Modeling for Age at Menarches (Log-Transformed) in Females across 37 Countries when the Female Life Expectancy Variable Was Replaced by the Mean Consumption in Vegetable Calorie and the Woman Illiteracy Rate per Country

| | <i>Parameter Estimate</i> | <i>Standard Error</i> | <i>df</i> | <i>F Value</i> | <i>Pr (F)</i> |
|------------------------|---------------------------|-----------------------|-----------|----------------|---------------|
| Intercept | 12.532 | 109.561 | 1 | | |
| (1) Latitude | -0.285 | 6.709 | 1 | 6.694 | 0.0953 |
| (2) Population density | -0.660 | 10.408 | 1 | 0.147 | 0.7090 |
| (3) Hemisphere | 11.008 | 107.301 | 1 | 1.931 | 0.1921 |
| (4) GNP | -0.682 | 6.582 | 1 | 18.459 | 0.3513 |
| (5) Sampling bias | -0.164 | 2.011 | 1 | 0.109 | 0.7470 |
| (6) Vegetable calorie | -0.001 | 2.011 | 1 | 4.041 | 0.0696 |
| (7) Illiteracy rate | 0.475 | 6.530 | 1 | 11.954 | 0.0054 |
| (1):(2) | 0.839 | 0.331 | 1 | 0.343 | 0.5701 |
| (1):(3) | -0.722 | 7.190 | 1 | 0.021 | 0.8883 |
| (1):(4) | -0.044 | 0.003 | 1 | 2.171 | 0.1687 |
| (1):(5) | 0.055 | 0.044 | 1 | 1.635 | 0.2274 |
| (1):(6) | <0.001 | <0.001 | 1 | 0.072 | 0.7936 |
| (1):(7) | 0.055 | 0.444 | 1 | 5.006 | 0.0869 |
| (2):(3) | -0.965 | 9.682 | 1 | 0.260 | 0.6200 |
| (2):(4) | -0.008 | 0.027 | 1 | 1.235 | 0.2902 |
| (2):(5) | -0.070 | 0.712 | 1 | 0.037 | 0.8514 |
| (2):(6) | <-0.001 | <0.001 | 1 | 1.390 | 0.2634 |
| (2):(7) | -0.015 | 0.013 | 1 | 0.921 | 0.3577 |
| (3):(4) | -0.965 | 9.682 | 1 | 2.538 | 0.1395 |
| (3):(5) | — | — | — | — | — |
| (3):(6) | <0.001 | 0.002 | 1 | 0.127 | 0.7286 |
| (3):(7) | -0.673 | 6.504 | 1 | 0.010 | 0.9231 |
| (4):(5) | — | — | — | — | — |
| (4):(6) | <0.001 | 0.002 | 1 | 0.014 | 0.9092 |
| (4):(7) | -0.007 | 0.012 | 1 | 0.490 | 0.4989 |
| (5):(6) | — | — | — | — | — |
| (5):(7) | — | — | — | — | — |
| (6):(7) | <0.001 | <0.001 | 1 | 0.595 | 0.4569 |

Note: Abbreviations are as in Tables 2 and 3. Statistics in bold characters correspond to variable(s) retained in minimal models.

menopause (e.g., Walsh 1978; Kirchengast 1992). More frequently, a significant positive relationship has been found between age at menopause and age at first birth (Stanford et al. 1987; Kirchengast 1992; Do et al. 1998). It is possible in our case that the negative relationship between fertility and age at menopause is in fact due to a positive relationship between the age at first birth and mean fertility (i.e., women with a high number of children are likely to have also started their reproductive life earlier). Although further information would be necessary to address this point, these findings support the idea that individual reproductive history influences the onset of the menopause.

Table 5. General Linear Modeling for Age at Menopause in Females across 21 Countries

| | <i>Parameter Estimate</i> | <i>Standard Error</i> | <i>df</i> | <i>F Value</i> | <i>Pr (F)</i> |
|---------------------|---------------------------|-----------------------|-----------|----------------|---------------|
| Intercept | 24.840 | 114.685 | 1 | | |
| (1) Latitude | -4.570 | 28.444 | 1 | 0.163 | 0.6979 |
| (2) Fertility | -23.323 | 54.067 | 1 | 15.685 | 0.0055 |
| (3) Age at menarche | 9.568 | 44.484 | 1 | 2.601 | 0.1509 |
| (4) Sampling bias | -10.091 | 38.216 | 1 | 2.020 | 0.1982 |
| (1):(2) | -1.664 | 1.889 | 1 | 0.680 | 0.4367 |
| (1):(3) | 2.165 | 11.190 | 1 | 0.188 | 0.6776 |
| (1):(4) | 0.158 | 0.920 | 1 | 0.248 | 0.6334 |
| (2):(3) | 8.207 | 19.306 | 1 | 0.001 | 0.9749 |
| (2):(4) | -0.025 | 3.833 | 1 | 0.249 | 0.6329 |
| (3):(4) | 3.859 | 14.605 | 1 | 0.292 | 0.6057 |
| (3):(5) | 0.187 | 0.151 | 1 | 1.537 | 0.2550 |

Note: Abbreviations are as in Tables 2 and 3. The variable number 4, i.e., sampling bias, means that values of age at menopause have been categorized into mean values (coded 1) or median values (coded 2) in order to check for the influence of this sampling artefact on calculation. The introduction into modeling of the interaction term (3):(5) means that values of age at menarches (coded 3) have been corrected for by the influence (coded 5) of being mean or median values (see Table 2 for further details). Statistics in bold characters correspond to variable(s) retained in minimal models.

In conclusion, the factors explaining interpopulation differences of age variability in menarche and menopause in human females are poorly known. Our present results demonstrate that on a large spatial scale different factors may be responsible for the observed variation in age at menarche, whereas a single factor may explain the age at menopause. In fact, age at menarche seems to be closely related to extrinsic factors such as living conditions and, especially, the energy balance allocated to individuals. Conversely, age at menopause appears to be more sensitive to intrinsic parameters, such as the reproductive history of individuals. Most of the work in this branch of science has focused on studies of the variation of events related to reproduction within populations. Important work remains to be done in specifying the patterns that determine human reproductive traits; such studies on a larger scale may offer some alternative interesting issues.

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