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Rapid response team and hospital mortality in hospitalized patients

Boris Jung^{1,2†}, Aurelien Daurat^{1†}, Audrey De Jong^{1,2}, Gerald Chanques^{1,2}, Martin Mahul¹, Marion Monnin¹, Nicolas Molinari³ and Samir Jaber^{1,2*}

Abstract

Purpose: Although rapid response systems are known to reduce in-hospital cardiac arrest rate, their effect on mortality remains debated. The present study aimed to evaluate the effect of implementing an intensivist-led rapid response team (RRT) on mortality in hospitalized patients.

Methods: An implementation of an intervention and a comparison with retrospective data analysis were performed in the four hospitals of Montpellier regional healthcare centre, in France. An intensivist-led RRT was implemented on a 24/7 basis along with educational modules, publicity and bedside simulation-based training in only one of the four hospitals from January 2012 to June 2012. A single activation criterion (heart rate below 40/min or above 140/min, systolic blood pressure below 80 mmHg, cardiac arrest, respiratory rate below 8/min or above 30/min, pulse oximetry below 90 % with O₂ above 6 l/min, respiratory distress in a tracheotomised patient, respiratory arrest, coma or sudden change in level of consciousness, seizure) allowed any caregiver to directly contact the RRT using a dedicated cell phone number. Patients over 18 years admitted for more than 24 h in the medical-surgical wards from July 2010 to December 2011 (pre-RRT period) and from July 2012 to December 2013 (RRT period) were included. The main outcome was unexpected mortality. Analyses of data from one RRT hospital and three control hospitals (no RRT hospital) were performed.

Results: RRT implementation was associated with a decrease in unexpected mortality rate in the hospital that implemented RRT (from 21.9 to 17.4 per 1000 discharges; $p = 0.002$). Reduction in unexpected mortality associated with RRT implementation could be estimated at 1.5 lives saved per week in the RRT hospital. In the three other hospitals, mortality rate was not significantly modified (from 19.5 to 19.9 per 1000 discharges; $p = 0.69$). Overall mortality decreased from 39.6 to 34.6 per 1000 discharges between the pre-RRT and RRT period in the RRT hospital ($p = 0.012$), but did not significantly change in the other hospitals. Patients in the RRT hospital were more frequently admitted to the intensive care unit (ICU) during the RRT period (45.8 vs 52.9 per 1000; $p = 0.002$), and their sequential organ failure assessment (SOFA) score upon ICU admission significantly decreased from 7 (4–10) to 5 (2–9); $p < 0.001$.

Conclusions: In the present retrospective study, implementation of an intensivist-led RRT along with educational modules, publicity and bedside simulation-based training was associated with a significant decrease in unexpected and overall mortality of inpatients.

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Take home message: The implementation of an intensivist-led rapid response team was associated with a decrease in both overall and unexpected mortality in a study performed in a French teaching hospital.

[†]B. Jung and A. Daurat contributed equally to the study.

Keywords: Medical emergency team, Rapid response team, Cardiac arrest, Patient safety

Introduction

In Europe, in-hospital cardiac arrest rate ranges from 1 to 5 per 1000 hospital discharges [1]. Traditionally, when a member of hospital staff witnesses a cardiac arrest, they call the cardiac arrest team to attend to this medical emergency. A rapid response system (RRS) is a hospital-based system designed to allow any staff member to alert other staff for help when a patient's vital signs have fallen outside set criteria. An intensivist-led rapid response team (RRT) is one in which an intensive care unit (ICU) doctor is a member of the response to the alert. The purpose of RRSs, which have mainly been developed in Northern America, Australia and Scandinavia, is to identify high-risk hospital patients early so that serious adverse events can be prevented and outcome improved [2, 3]. Two previous meta-analyses concluded that RRSs were associated with a reduction in cardiac arrest but that the effect on mortality was doubtful [4, 5]. The lack of a definite impact on outcome may have been due to several methodological issues from insufficient utilization to delays to activate the RRS [6]. Despite this lack of evidence, RRSs have been widely adopted in Australia and North America during the last decade [7, 8]. Implementation in Europe remains rare and this study is one of the first such initiatives published in mainland Europe [9]. Our regional healthcare centre comprises four different hospitals. Each hospital is equipped with an independent ICU and the distances between them require an ambulance if a patient needs to be transferred.

The aim of the study was to assess whether the implementation of an intensivist-led RRT available only in one hospital would be associated with a decrease in the incidence of cardiac arrest and an increase in overall and unexpected hospital mortality.

Methods

The Comité d'Organisation et de Gestion de l'Anesthésie Réanimation (COGAR), a university hospital ethics committee in Montpellier, France, reviewed and approved the study. The need for informed consent was waived on the basis that the intervention was a quality improvement initiative.

Study design

We studied the effect of an RRT implemented in one hospital and compared that to retrospective data from three control hospitals. We included all consecutive adult patients admitted to the regional healthcare centre for at least 24 h between July 2010 and December 2013.

We divided the cohort into three temporal cohorts: the pre-RRT period (18 months, from July 2010 to December 2011), the implementation period (6 months, from January 2012 to June 2012) and the RRT period (18 months from July 2012 to December 2013).

In order to examine in-hospital outcome among patients most likely to benefit from an RRT intervention, we excluded patients hospitalized in the long-term geriatric care facility, those receiving maternity care, patients in the rehabilitation day-care facility and mental health services [10]. In these excluded units, specific emergency responses exist and involve ambulances with emergency physicians and none of these patients were hospitalized in an area within reach of the RRT initiative team.

The reporting of this study follows the Strengthening the Reporting of Observational Studies in Epidemiology (STROBE) guidelines [11] and the REporting of studies Conducted using Observational Routinely-collected health Data (RECORD) Statement [12]. The data were collected using dedicated forms designed according to the in-hospital resuscitation Utstein style [13] in the post-implementation period in the RRT hospital and using the hospital database (Programme de Médicalisation des Systèmes d'Information, PMSI, Montpellier University Hospital) for the pre-implementation period in the RRT hospital. The hospital database was used for both pre- and post-implementation periods in the three non-RRT hospitals.

Setting

As routinely observed in French teaching hospitals, hospital wards had a nurse/patient ratio ranging from 1:12 in the daytime (6.30 a.m.–9 p.m.) to 1:20 at night (9 p.m.–6.30 a.m.) and the same ratio of assistant nurses. All ICUs were organized according to a closed format system and covered by a full team of attendings, fellows and residents during daytime (8 a.m.–6 p.m.) and one attending and one resident in house during night-time. Respiratory therapists, physician assistants and nurse practitioners are professions that do not exist in France. During the study period, there was no change in admission criteria, nursing or medical care plan or equipment. The RRT hospital case mix was mostly represented by abdominal and transplant surgery, liver and gastroenterology clinics, internal medicine, dermatology and haematology wards while the non-RRT hospitals provided all other specialities including cardiology and cardiac surgery. The coronary care unit was considered as a general ward, as was a step-down unit that did not belong to an ICU

department. The non-RRT hospitals were chosen by convenience because of the location of health centre.

Implementation

In the pre-RRT period, all ICU teams responded to codes and triage calls using a traditional pyramid set up, involving the bedside nurse paging the resident then the fellow in charge who escalated to the attending on call. Either the fellow or the attending would eventually call the ICU nurse station to request a triage or a code blue response. No record of activity was obtained in that period. A code blue was defined as any patient with an unexpected cardiac or respiratory arrest requiring resuscitation [14].

In the implementation period, the RRT was instigated as a plan-do-study-act project [15, 16] in only one of the hospitals as a pilot. A multidisciplinary group composed of three attending intensivists (BJ, GC, SJ), the St Eloi ICU head nurse and an ICU fellow developed an activation criteria list adapted from the existing literature (Table 1) [17, 18].

A single criterion on that list allowed any caregiver to directly contact the RRT using a dedicated phone number (Fig. 1). The RRT comprised an ICU resident and either an ICU fellow or an attending. An ICU nurse could be part of the team if requested by the attending. No extra funding or caregiver hiring was associated with this pilot initiative. The RRT was equipped with a crash cart that included a portable ventilator, a monitor, a capnometer and oxygen [19]. The RRT was expected to reach the scene within 5 min in case of a code blue response and within 20 min in other situations. The RRT was expected to communicate with the ward team using the SBAR (situation, background, assessment, recommendation) scheme for communication and could provide any procedure, drugs prescription, facilitate an ICU transfer or discuss an end of life issue [20]. During a 6-month period, the RRT activation criteria were presented to the medical and nursing teams that were in the St Eloi ICU perimeter. Presentation included displaying posters, bedside simulation-based training courses using manikins for ward residents, doctors and nurses, practical educational sessions and information through the local hospital newspaper.

Data collection and outcome assessment

The study is a retrospective evaluation of the RRT intervention in one hospital (RRT hospital) complemented with data from three other hospitals (non-RRT hospitals). RRT intervention data and ICU data (demographics and outcome) in the RRT hospital were retrospectively analysed. Non-RRT ICU data (demographics and outcome) and inpatient data in all hospitals were retrospectively collected from the hospital billing codes (Programme de Médicalisation des Systèmes d'Information, PMSI).

Patient illness acuity or case-mix index across the different hospitals and during the study was categorized on a scale ranging from 1 (mild) to 4 (severe) according to the recommendations of the French Technical Agency on Hospital Information [21]. This index takes into account the main diagnosis, length of hospital stay and complications that might occur during the hospital stay. A senior hospital database manager, blinded to the study design and the scientific rationale, provided the case-mix index and the inpatient mortality rates. A surgical case was defined as one in which any surgery took place during the hospital stay.

The main endpoint of the study was the unexpected mortality rate defined as hospital-wide non-DNR and non-palliative care unplanned death per 1000 discharges as previously reported [22, 23]. The main endpoint was chosen according to guidelines for medical emergency teams [14]. DNR/palliative care order could be made during the hospital stay by the attending on service or by an intensivist in close collaboration with the ward doctors. Secondary endpoints were overall mortality, cardiac arrest rate per 1000 discharges occurring outside the ICU, unplanned ICU admission rate per 1000 discharges from medical and surgical wards (available only in the RRT hospital ICU), do not resuscitate patient deaths, and hospital length of stay. A cardiac arrest was defined as a cardiac arrest that was resuscitated.

Statistical analysis

On the basis of unpublished personal data of the RRT hospital (about 15,000 admissions per year, with an unexpected mortality of 2.1 per 1000 discharges), we estimated the number of subjects required for this study using the Arcsin approximation method. Assuming a relative reduction of 20 % in unexpected mortality with a 5 % α risk and a β risk of 90 %, we estimated that 22,023 subjects per period were needed. Consequently, with a yearly basis of about 15,000 admissions in the medical-surgical wards of the St Eloi Hospital, we set the observation periods duration to 18 months each. Normally distributed quantitative data were described as mean \pm standard deviation (SD) and compared using Student's *t* test. If not normally distributed, quantitative data were expressed as median and interquartile range (IQR) and compared using the Mann–Whitney test. Qualitative data were expressed as number (rate) and compared using the Chi square test with Yates' correction as appropriate. Endpoints (monthly unexpected and overall mortality, cardiac arrest rate) were compared pre-RRT implementation (18-month period) and post-RRT implementation (18-month period) in the RRT hospital using a Mann–Whitney test. To limit the risk of reporting outcome results related to natural evolution or chance, we also compared the same endpoints

Table 1 Demographic and outcome characteristics of the patients during the three study periods according to the hospital

a	RRT hospital (n = 43,605)																								
	Hospital 1 (n = 49,917)					Hospital 2 (n = 37,557)					Hospital 3 (n = 29,992)														
	Pre	Per	Post	p		Pre	Per	Post	p		Pre	Per	Post	p											
	(n = 18,072) (n = 6461) (n = 19,073)					(n = 22,158) (n = 7077) (n = 20,682)					(n = 14,837) (n = 5843) (n = 16,877)					(n = 13,019) (n = 4439) (n = 12,534)									
Age (years)	59 ± 19	60 ± 18	59 ± 18	0.35	55 ± 19	56 ± 19	57 ± 19	<0.001	55 ± 18	55 ± 18	56 ± 18	0.032	64 ± 15	65 ± 15	65 ± 15	<0.001									
Men	9599 (53)	3333 (52)	10,194 (53)	0.52	11,028 (50)	3500 (49)	10,387 (50)	0.35	8066 (54)	3053 (52)	8978 (53)	0.039	8612 (66)	2938 (66)	8338 (67)	0.54									
Severity class																									
I/II	13,030 (72)	4721 (73)	13,565 (71)	0.038	19,154 (86)	6009 (85)	17,241 (83)	<0.001	12,330 (83)	4922 (84)	14,175 (84)	0.034	11,398 (88)	3884 (87)	11,132 (89)	0.0019									
III/IV	5042 (28)	1740 (27)	5507 (29)		3004 (14)	1068 (15)	3441 (17)		2507 (17)	921 (16)	2702 (16)		1621 (12)	555 (13)	1402 (11)										
Admission																									
Surgical	3565 (20)	1304 (20)	4082 (21)	<0.001	10,295 (46)	3217 (45)	9160 (44)	<0.001	5771 (39)	2429 (42)	7132 (42)	<0.001	3270 (25)	1190 (27)	3191 (27)	0.54									
Medical	14,507 (80)	5157 (80)	14,991 (79)		11,863 (54)	3860 (55)	11,522 (56)		9066 (61)	3414 (58)	9745 (58)		9749 (75)	3249 (73)	9343 (75)										
Unexpected mortality	396 (21.9)	150 (23.2)	332 (17.4)	0.002	317 (14.3)	96 (13.6)	318 (15.4)	0.38	370 (24.9)	137 (23.4)	379 (22.5)	0.16	288 (22.1)	119 (26.8)	298 (23.8)	0.40									
Non-ICU cardiac arrests	48 (2.6)	13 (2.0)	34 (1.8)	0.093	77 (3.5)	19 (2.7)	95 (4.6)	0.080	49 (3.3)	8 (1.4)	35 (2.1)	0.044	133 (10.2)	55 (12.4)	135 (10.8)	0.71									
ICU admis-sion	827 (45.7)	312 (48.3)	1008 (52.8)	0.002	1131 (51.0)	361 (51.0)	1143 (55.3)	0.054	893 (60.2)	323 (55.3)	972 (55.6)	0.34	1652 (126.9)	588 (132.5)	1529 (122.0)	0.24									
Death with DNR order	319 (17.7)	104 (16.1)	328 (16.4)	0.6	52 (2.3)	15 (2.1)	62 (3.0)	0.22	55 (3.7)	21 (3.6)	47 (2.8)	0.18	89 (6.8)	24 (5.4)	65 (5.2)	0.10									
Overall mor-tality	715 (39.6)	254 (39.3)	660 (34.6)	0.012	369 (16.7)	111 (15.7)	380 (18.4)	0.19	425 (28.6)	158 (27.0)	426 (25.2)	0.066	377 (29.0)	143 (32.2)	363 (29.0)	0.97									
Hospital length of stay (days)	5 (2–10)	4 (2–10)	5 (2–10)	0.09	4 (2–8)	4 (2–8)	4 (2–8)	<0.001	3 (2–7)	3 (2–7)	3 (2–6)	<0.001	4 (2–8)	4 (2–8)	4 (2–8)	0.36									
b	RRT hospital (n = 43,605)										Other hospitals (n = 117,466)														
	Pre	Per	Post	p		Pre	Post				Pre	Per	Post	p											
	(n = 18,072)					(n = 6461)					(n = 19,073)					(n = 50,014)					(n = 50,993)				
	Age (years)	59 ± 19	60 ± 18	59 ± 18	0.35	55 ± 19	56 ± 18	57 ± 19	<0.001	55 ± 18	55 ± 18	56 ± 18	0.032	64 ± 15	65 ± 15	65 ± 15	<0.001								
	Men	9599 (53)	3333 (52)	10,194 (53)	0.52	11,028 (50)	3500 (49)	10,387 (50)	0.35	8066 (54)	3053 (52)	8978 (53)	0.039	8612 (66)	2938 (66)	8338 (67)	0.54								
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Medical	14,507 (80)	5157 (80)	14,991 (79)		11,863 (54)	3860 (55)	11,522 (56)		9066 (61)	3414 (58)	9745 (58)		9749 (75)	3249 (73)	9343 (75)										
Unexpected mortality	396 (21.9)	150 (23.2)	332 (17.4)	0.002	317 (14.3)	96 (13.6)	318 (15.4)	0.38	370 (24.9)	137 (23.4)	379 (22.5)	0.16	288 (22.1)	119 (26.8)	298 (23.8)	0.40									
Non-ICU cardiac arrests	48 (2.6)	13 (2.0)	34 (1.8)	0.093	77 (3.5)	19 (2.7)	95 (4.6)	0.080	49 (3.3)	8 (1.4)	35 (2.1)	0.044	133 (10.2)	55 (12.4)	135 (10.8)	0.71									
ICU admis-sion	827 (45.7)	312 (48.3)	1008 (52.8)	0.002	1131 (51.0)	361 (51.0)	1143 (55.3)	0.054	893 (60.2)	323 (55.3)	972 (55.6)	0.34	1652 (126.9)	588 (132.5)	1529 (122.0)	0.24									
Death with DNR order	319 (17.7)	104 (16.1)	328 (16.4)	0.6	52 (2.3)	15 (2.1)	62 (3.0)	0.22	55 (3.7)	21 (3.6)	47 (2.8)	0.18	89 (6.8)	24 (5.4)	65 (5.2)	0.10									
Overall mor-tality	715 (39.6)	254 (39.3)	660 (34.6)	0.012	369 (16.7)	111 (15.7)	380 (18.4)	0.19	425 (28.6)	158 (27.0)	426 (25.2)	0.066	377 (29.0)	143 (32.2)	363 (29.0)	0.97									
Hospital length of stay (days)	5 (2–10)	4 (2–10)	5 (2–10)	0.09	4 (2–8)	4 (2–8)	4 (2–8)	<0.001	3 (2–7)	3 (2–7)	3 (2–6)	<0.001	4 (2–8)	4 (2–8)	4 (2–8)	0.36									
b	RRT hospital (n = 43,605)										Other hospitals (n = 117,466)														
	Pre	Per	Post	p		Pre	Post				Pre	Per	Post	p											
	(n = 18,072)					(n = 6461)					(n = 19,073)					(n = 50,014)					(n = 50,993)				
	Age (years)	59 ± 19	60 ± 18	59 ± 18	0.35	55 ± 19	56 ± 18	57 ± 19	<0.001	55 ± 18	55 ± 18	56 ± 18	0.032	64 ± 15	65 ± 15	65 ± 15	<0.001								
	Men	9599 (53)	3333 (52)	10,194 (53)	0.52	11,028 (50)	3500 (49)	10,387 (50)	0.35	8066 (54)	3053 (52)	8978 (53)	0.039	8612 (66)	2938 (66)	8338 (67)	0.54								
Severity class																									
I/II	13,030 (72)	4721 (73)	13,565 (71)	0.038	19,154 (86)	6009 (85)	17,241 (83)	<0.001	12,330 (83)	4922 (84)	14,175 (84)	0.034	11,398 (88)	3884 (87)	11,132 (89)	0.0019									
III/IV	5042 (28)	1740 (27)	5507 (29)		3004 (14)	1068 (15)	3441 (17)		2507 (17)	921 (16)	2702 (16)		1621 (12)	555 (13)	1402 (11)										
Admission																									
Surgical	3565 (20)	1304 (20)	4082 (21)	<0.001	10,295 (46)	3217 (45)	9160 (44)	<0.001	5771 (39)	2429 (42)	7132 (42)	<0.001	3270 (25)	1190 (27)	3191 (27)	0.54									
Medical	14,507 (80)	5157 (80)	14,991 (79)		11,863 (54)	3860 (55)	11,522 (56)		9066 (61)	3414 (58)	9745 (58)		9749 (75)	3249 (73)	9343 (75)										
Unexpected mortality	396 (21.9)	150 (23.2)	332 (17.4)	0.002	317 (14.3)	96 (13.6)	318 (15.4)	0.38	370 (24.9)	137 (23.4)	379 (22.5)	0.16	288 (22.1)	119 (26.8)	298 (23.8)	0.40									
Non-ICU cardiac arrests	48 (2.6)	13 (2.0)	34 (1.8)	0.093	77 (3.5)	19 (2.7)	95 (4.6)	0.080	49 (3.3)	8 (1.4)	35 (2.1)	0.044	133 (10.2)	55 (12.4)	135 (10.8)	0.71									
ICU admis-sion	827 (45.7)	312 (48.3)	1008 (52.8)	0.002	1131 (51.0)	361 (51.0)	1143 (5																		

Table 1 continued

b	RRT hospital (n = 43,605)				Other hospitals (n = 117,466)			
	Pre	Per	Post	p	Pre	Per	Post	p
	(n = 18,072)	(n = 6461)	(n = 19,073)		(n = 50,014)	(n = 17,359)	(n = 50,093)	
Surgical	3565 (20)	1304 (20)	4082 (21)	<0.001	19,336 (39)	6836 (39)	19,483 (39)	0.45
Medical	14,507 (80)	5157 (80)	14,991 (79)		30,678 (61)	10,523 (61)	30,610 (61)	
Unexpected mortality	396 (2.2)	150 (2.3)	332 (1.7)	0.002	975 (1.9)	352 (2)	995 (2)	0.69
Non-ICU cardiac arrests	48 (0.26)	13 (0.20)	34 (0.18)	0.093	259 (0.52)	82 (0.47)	265 (0.53)	0.84
ICU admission	827 (4.6)	312 (4.8)	1008 (5.3)	0.002	3676 (7.4)	1272 (7.3)	3644 (7.3)	0.65
Death with DNR order	319 (1.8)	104 (1.6)	328 (1.6)	0.6	196 (0.9)	60 (0.4)	174 (0.4)	0.27
Overall mortality	715 (4.0)	254 (3.9)	660 (3.4)	0.012	1171 (2.3)	412 (2.4)	1169 (2.3)	0.95
Hospital length of stay (days)	5 (2–10)	4 (2–10)	5 (2–10)	0.09	4 (2–8)	4 (2–8)	4 (2–8)	0.10

Data are presented as number and percentage or median and quartiles; p value is between pre- and post-intervention periods
Pre pre-intervention period, Per pre-implementation period, Post post-intervention period, DNR do not resuscitate, RRT rapid response team

between the two 18-month periods in the three non-RRT hospitals [24]. Because the case mix was different among the four hospitals, an adjustment was made on the basis of the average age, gender and the average case-mix index using a linear and mixed regression model taking into account the month of hospitalization. A Poisson regression was also performed taking into account age, gender and severity index and the number of patients per month as offset variable to estimate the number of patients needed to be admitted in wards to save one unexpected death. Statistical analysis was performed using R software (version 3.0.2).

Results
Demographics

During the whole study period, a total of 161,071 patients were admitted for 24 h or more in the medical-surgical wards of the four healthcare centre hospitals. There were 68,086 patients during the pre-RRT period, and 69,165 in the RRT period. Demographic characteristics of patients during the two study periods according to hospital and period are summarized in Table 1.

Unexpected mortality

The unexpected mortality rate for 1000 discharges significantly decreased from 21.9 to 17.4 between the pre-RRT and post-RRT periods ($p = 0.002$).

After adjustment, the RR was calculated as equal to 0.77 (95 % CI 0.61–0.99) (Table 1). Assuming this relative reduction the number of patients hospitalized in wards needed to save one life was 225 (95 % CI 130–3788), $p = 0.04$. With an average of about 18,000 admission/year in the RRT hospital and on average one RRT activation per day, our initiative was associated with 1.5 lives saved a week. In the meantime no significant change was noted in mortality rates in the three hospitals that did not implement RRT (Fig. 2). Unexpected hospital mortality (including ICU mortality) following sepsis decreased in the RRT hospital [109/2571 (4.2 %) before RRT implementation vs 82/2638 (3.1 %) after RRT implementation, $p = 0.03$] but not in the three others hospitals with no RRT. In the RRT hospital, among the 142 septic ward patients for whom an RRT was called, 49 were admitted to the ICU and 11 of them (22 %) ultimately died. No data was available before RRT implementation. Data on the distinction between planned vs unplanned ICU admissions were collected only in the RRT hospital. The SOFA upon admission for unplanned admissions was 7 (4–10) before RRT implementation and 5 (2–9), $p < 0.01$ after RRT implementation. The ICU mortality rate for these unplanned patients was 121/690 (17.5 %) before RRT implementation and 150/820 (18.3 %; $p = 0.71$) after RRT implementation.

a

CLINICAL WARNING SIGNS

CIRCULATION

- ✓ Cardiac Arrest
- ✓ Heart rate below 40/min or above 140/min
- ✓ Systolic Blood Pressure below 80 mmHg


RESPIRATORY

- ✓ Respiratory Arrest
- ✓ Acute Respiratory Failure (respiratory rate below 8/min or above 30/min)
- ✓ Pulse Oxymetry (SpO₂) below 90% with O₂ above 6l/min
- ✓ Respiratory Distress in a tracheotomised patient


CONSCIOUSNESS

- ✓ Coma or sudden change in level of consciousness
- ✓ Seizure

b



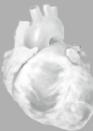
Saint Eloi Anesthesiology and Intensive Care Unit (Pr Samir Jaber)
Dr Boris Jung (Intensive Care Unit Director)



MET

Medical Emergency Team

You are a caregiver and are concerned about the sudden change in the clinical status of one of your patient or you need an ICU consult right away?



Clinical Warning Signs

CIRCULATION



- * Cardiac arrest
- * Heart Rate < 40 bpm or > 140 bpm
- * Systolic Blood Pressure < 80 mmHg

RESPIRATORY

- * Respiratory arrest
- * Acute Respiratory Failure (respiratory rate < 8/min or > 30/min)
- * Oxygen Saturation (SpO₂) < 90% with O₂ > 6l/min
- * Respiratory distress in a tracheotomized patient

CONSCIOUSNESS

- * Coma or sudden changes in level of consciousness
- * Seizure



Call 24/7 : 3 72 72
or Cell : 1 91 98



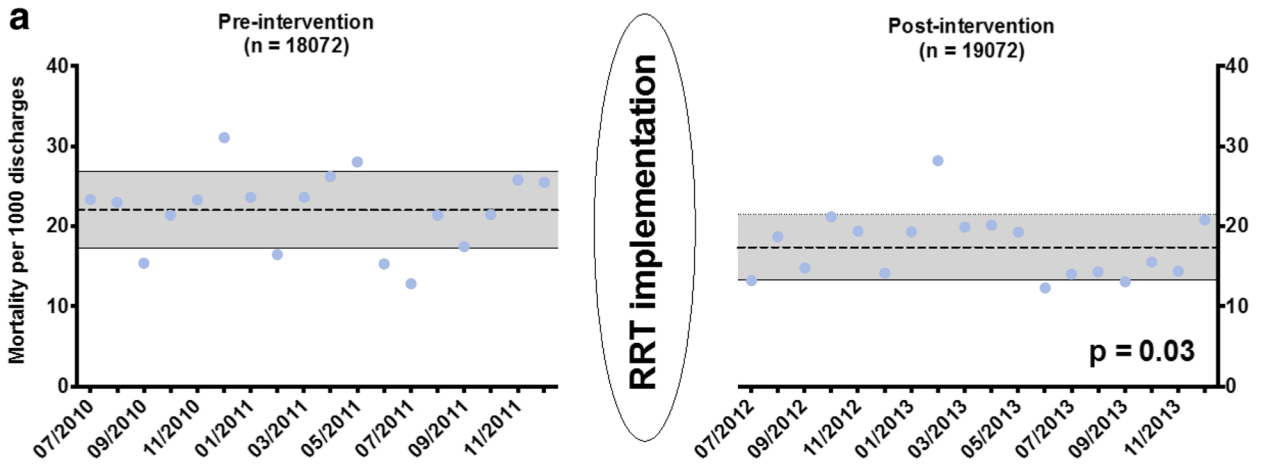


Fig. 1 a Activation criteria. A single criterion allowed any caregiver to directly reach the rapid response team (RRT) using a dedicated cell phone number. Because of their high-risk profile, the RRT could be activated following a subjective assessment in tracheotomised patients. **b** Poster reminding staff of the activation criteria and phone numbers of the medical emergency team

RRT hospital



Other hospitals

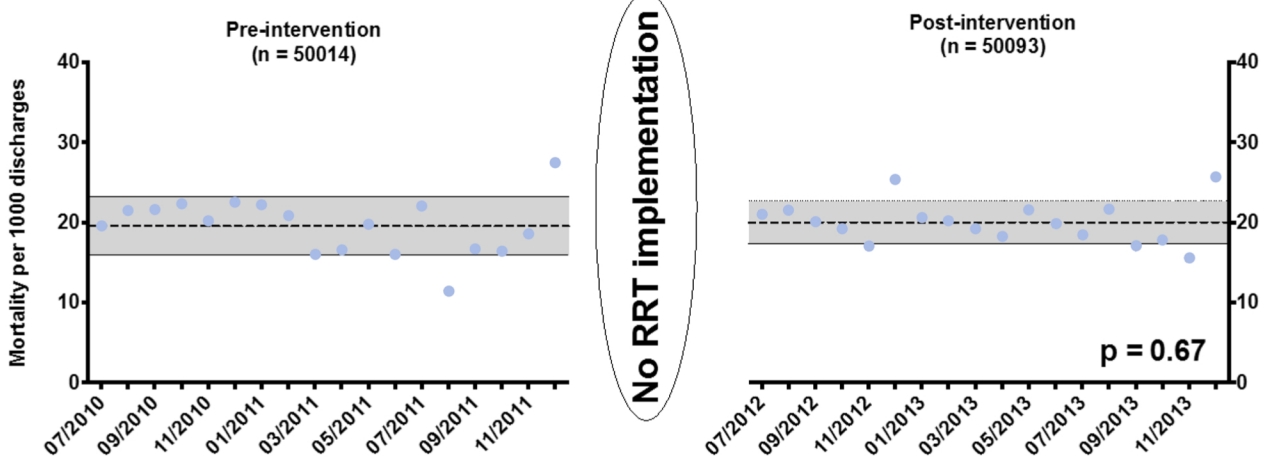


Fig. 2 Evolution of unexpected death rate per 1000 discharges by month in the RRT hospital (**a**), and in the three other hospitals (**b**). Dotted lines represent the mean rate per month. Grey rectangles represent the standard deviations

Overall mortality and cardiac arrests

Overall mortality decreased from 39.6 to 34.6 per 1000 discharges between the pre-RRT and RRT period in the RRT hospital ($P = 0.012$), but did not significantly change in the other hospitals (Table 1). We observed a decrease in cardiac arrest in the intervention period but this was not significant (2.6 vs 1.8 per 1000 admissions; $p = 0.07$) so it could be due to chance alone. No trend in cardiac arrest incidence rate was observed in the three other hospitals (5.2 vs 5.3 per 1000 discharges, respectively in pre-RRT and RRT period, $p = 0.84$). Rates of deaths with DNR orders and hospital length of stay did not change significantly between the two periods in any hospital (Table 1).

RRT interventions

During the 18-month RRT period, a total of 564 RRT interventions were carried out (i.e. 29.6 interventions per 1000 discharges). Median time of arrival was 5 min (5–10), and main activation criteria was acute hypoxaemia defined by $\text{SpO}_2 < 90\%$ (Table 2). The most common initial diagnosis was sepsis that accounted for 29 % of the interventions (Table 2). Patients hospitalized in the RRT hospital during the RRT-period were more likely to be transferred to the ICU during their stay (pre-MET period 45.8 vs 52.9 per 1000 discharges in RRT period; $p = 0.002$), whereas rates of ICU admission did not statistically change in other hospitals (73.5 vs 72.7 per 1000

discharges; $p = 0.66$). In patients admitted to the ICU, median age significantly increased in the RRT hospital during the RRT period [59 (47–69) vs 61 (50–71); $p = 0.03$]. In these patients, SOFA score was lower during the RRT period compared with the pre-RRT period: 7 (4–10) vs 5 (2–9); $p < 0.001$. When planned admissions were not analysed, the SOFA score still decreased after the RRT implementation [7 (4–10) vs 5 (2–9); $p < 0.01$]. In the RRT ICU, there was no overall significant change in SAPSII, ICU mortality, ICU length of stay and mechanical ventilation duration between the two study periods (Table 3).

Discussion

In this study, RRT implementation was associated with a 20.6 % drop in unexpected mortality. In the meantime, the rate of ICU transfer from the wards increased and the organ failure score (SOFA) upon ICU admission of these patients decreased. None of these changes were noted in the three other hospitals that did not implement RRT. There was no significant difference between the two periods for cardiac arrest rate in any of the hospitals.

Until now, the effect of RRS on mortality has been debated with several discordant studies [25]. Two randomised clinical trials examined the impact of RRT on patient outcome [26, 27]. In the large MERIT trial, 120,000 patients were enrolled in 23 hospitals in Australia but the mean calling rate in the RRT group was perhaps too low (8.7 calls per 1000 discharges) to have an impact on outcome. Moreover, as the RRT was part of a national initiative and was advertised in the media, the authors suspected a contamination effect of the control group. A second cluster-ward randomised trial performed in a British hospital evaluated a nurse-driven RRT and reported a positive effect on patient mortality [26], but the prospective cluster-ward design might have generated a Hawthorne effect. The meta-analysis by Chan et al. [28] reported a reduction in cardiac arrest rate but no significant effect on mortality [4]. More recently, the updated review by Winters et al. suggested a positive impact on patient outcome; however, statistical significance was not reached [5].

The RRT initiative in the present study meets quality criteria [14]: the “dose” of RRT delivered was 29.6 per 1000 patients, which is above the minimal rate of 25 per 1000 recommended [6], and median intervention time was short (5 min). The ward nurse staffing followed the French recommendations and the case-mix index did not change between the study periods. We can, however, not speculate to what extent the RRT initiative would have been associated with an improved outcome if the nurse–patient ratio had been higher.

Table 2 Main activation criteria, primary actions, immediate triage decisions and bedside diagnosis after RRT intervention

RRT activations	564 (100)
RRT main activation criteria	
SpO ₂ <90 %	97 (17)
SBP <80 mmHg	79 (14)
Altered mental status	76 (14)
Respiratory rate >30 c/min	54 (10)
Dyspnoea	45 (8)
Unspecific clinical concern	36 (6)
Heart rate >140 bpm	35 (6)
Cardiac arrest	27 (5)
Haemorrhage	21 (4)
Seizure	13 (2)
Heart rate <40 bpm	4 (1)
Others	74 (13)
Position of activator calling RRT	
Nurse	167 (30)
Resident	273 (48)
Fellow or attending	105 (19)
Non medical staff	19 (3)
Main primary action performed by RRT	
Crystalloid infusion	123 (22)
Antibiotics	43 (8)
Peripheral line	39 (7)
Intubation	36 (6)
Colloid infusion	36 (6)
Diuretics	36 (6)
Strategy advice only	32 (6)
Transfusion	31 (5)
Compressions for acute bleeding	25 (4)
Nebulizer treatment	22 (4)
Vasopressor	21 (4)
Antiarrhythmic treatment	16 (3)
Non-invasive ventilation	12 (2)
Defibrillation	10 (2)
Flumazenil	10 (2)
Central venous access	6 (1)
Analgesics	5 (1)
Antiepileptic drugs	5 (1)
Others	47 (8)
Immediate triage	
Patient was not an ICU candidate	348 (62)
Too sick to benefit	66 (12)
Too well to benefit	282 (50)
ICU admission	187 (33)
Death after resuscitation	19 (4)
Missing data	10 (2)
Initial diagnosis	
Sepsis	142 (25)
Severe sepsis/sepsis shock	121 (21)

Table 2 continued

RRT activations	564 (100)
Non severe sepsis	11 (2)
Immunodepression-related sepsis	10 (2)
Haemorrhage/haemorrhagic shock	45 (8)
Hypovolaemia	11 (2)
Acute heart failure/cardiogenic shock	52 (9)
Arrhythmia	12 (2)
Pulmonary embolism	11 (2)
Acute on chronic liver failure	38 (7)
Acute pancreatitis	11 (2)
Stroke	9 (2)
Seizure	17 (3)
Altered mentality	4 (1)
Anaphylaxis	15 (3)
COPD exacerbation	15 (3)
Acute renal failure	8 (1)
Drug overdose	7 (1)
Others	91 (16)
Missing data	76 (13)

Data are expressed as number and percentages of total RRT intervention (n = 564)

COPD chronic obstructive pulmonary disease, ICU intensive care unit, RRT rapid response team, SpO2 oxyhaemoglobin saturation measured by pulse oximetry, SBP systolic blood pressure

The team was led by an ICU physician which might have contributed to a greater efficiency [28]. Our initiative concerned one area of the regional healthcare centre while other areas depended on other ICUs with no RRT. To avoid daytime rounds and disruption of tasks, the 1–2 activations per 24 h were mostly provided by attendings assigned to non-clinical duty and by the in-house team on call during night-time. The RRT implementation was associated with a higher rate of early admission of less sick and older patients (Table 3), a double-edged sword effect linked to RRT systems [30, 31] but associated with a decrease in mortality at a minimal cost in the present study. Reduction of unexpected mortality associated with RRT implementation is partially related to a positive impact on outcome in patients diagnosed with sepsis. Our data are in line with the hypothesis that RRT might have taught caregivers to respond better to sepsis in patients admitted in wards but also provided a rapid ICU consultation to respond early to a patient presenting a progressing sepsis. This is also in line with extensive literature on early goal-directed therapy and the Surviving Sepsis Campaign that emphasize the need for rapid response in sepsis patients [32, 33].

Notably, RRT implementation was not associated with a statistically significant decrease in cardiac arrest rate in the present study possibly because of lack of power or because the rate was already low before the implementation. However, we observed a decrease in cardiac arrest rate in hospital 2 possibly because of a decrease in the patients’ severity in this hospital (Table 1a).

The present study has several strengths. First, we enrolled a large number of patients (i.e. 137,251) over a consecutive 3-year-period and we performed the implementation of an RRT intervention (only RRT hospital) and a comparison with retrospective data analysis (both RRT and non-RRT hospitals). However, as the case mix was different among the four hospitals, we adjusted for patient’s severity (based on the case-mix index), gender and age.

Several limitations can also be discussed. All planned and unplanned admissions were considered because this distinction was not available in the non-RRT hospitals. A single RRT activation parameter was used in this study, although aggregate weighted scoring systems may be more accurate but also less user-friendly [29]. We did not evaluate health-care-related costs but the cost of the intervention may have been minimized by the fact that the RRT team was part of the ICU routine care without hiring extra staff. DNR orders and cardiac arrest rate were collected from the hospital electronic data reports in the four hospitals and could have been preceded by RRT activation. No modifications of ICU protocols were made that could have influenced the outcome of patients admitted after RRT activation but the ICU team in the RRT hospital was motivated to provide the best service including early ICU admission for high-risk patients. It would have been of interest to report the call rate during a fourth period to check if the call rate was sustained after implementation but data this was not collected. Motivation may be viewed as an interpretation bias but is necessary in any RRT implementation. It took significant efforts to launch the RRT initiative and to train ICU staff but also to convince ward doctors and nurses to contact the RRT for high-risk patients without delay. A training and educational module including bedside simulation with a manikin, information using posters and advertisement was designed to help caregivers, and the nurses in particular, to recognize the listed activation criteria.

Conclusion

Intensivist-led RRT implementation with educational modules for nurses and doctors, bedside simulation-based training and publicity was associated with a significant decrease in unexpected and overall mortality of hospitalized patients.

Table 3 Characteristics and outcome of patients admitted to the ICU according to the hospital during the study periods

	RRT hospital ICU (n = 2148)				Hospital 1 ICU (n = 2635)				Hospital 2 ICU (n = 2188)				Hospital 3 ICU (n = 3769)			
	Pre (n = 827)	Per (n = 312)	Post (n = 1009)	p	Pre (n = 1131)	Per (n = 361)	Post (n = 1143)	p	Pre (n = 893)	Per (n = 323)	Post (n = 972)	p	Pre (n = 1652)	Per (n = 588)	Post (n = 1529)	p
ICU deaths	122 (14.7)	43 (13.8)	140 (13.9)	0.64	218	64	224	0.89	246	85	249	0.37	148	58	156	0.26
Age	59 (47–69)	61 (48–71)	61 (50–71)	0.032	55 (37–70)	56 (38–69)	55 (38–71)	0.68	60 (47–72)	62 (49–71)	60 (47–70)	0.69	66 (56–75)	67 (57–75)	66 (57–75)	0.12
SAPS II	36 (25–48)	35 (27–47)	36 (26–50)	0.40	35 (22–51)	34 (21–50)	35 (21–52)	0.93	38 (27–52)	39 (28–52)	39 (27–54)	0.39	25 (18–34)	25 (18–36)	25 (20–37)	0.016
ICU length of stay (days)	4 (2–9)	3 (2–9)	4 (2–8)	0.55	3 (1–9)	4 (2–11)	3 (1–9)	0.59	6 (2–15)	7 (3–19)	6 (2–17)	0.19	1 (1–2)	1 (1–3)	1 (1–3)	0.078

Pre pre-intervention period, Per per-implementation period, Post post-intervention period, DNR do not resuscitate

p value is between pre- and post-intervention periods

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Compliance with ethical standards

Conflicts of interest

Boris Jung reports personal fees from Merck (Whitehouse station, NJ) and Astellas (Tokyo, Japan) not related to the present study. Aurelien Daurat, Audrey De Jong, Martin Mahul, Marion Monnin, Gerald Chanques and Nicolas Molinari have nothing to disclose. Samir Jaber reports personal fees from Maquet, Draeger, Hamilton Medical and Fisher Paykel not related to the present study.

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